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A NEW SPECIES OF ORANGUTAN

Report

Morphometric, behavioral, and genomic evidence for a new orangutan species

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73 **Summary**

74 Six extant species of non-human great apes are currently recognized: Sumatran and Bornean orangutans,
75 eastern and western gorillas, and chimpanzees and bonobos [1]. However, large gaps remain in our
76 knowledge of fine-scale variation in hominoid morphology, behavior, and genetics, and aspects of great
77 ape taxonomy remain in flux. This is particularly true for orangutans (genus: *Pongo*), the only Asian
78 great apes, and phylogenetically our most distant relatives among extant hominids [1]. Designation of
79 Bornean and Sumatran orangutans, *P. pygmaeus* (Linnaeus 1760) and *P. abelii* (Lesson 1827), as distinct
80 species occurred in 2001 [1, 2]. Here, we show that an isolated population from Batang Toru, at the
81 southernmost range of extant Sumatran orangutans south of Lake Toba, is distinct from other northern
82 Sumatran and Bornean populations. By comparing cranio-mandibular and dental characters of an
83 orangutan killed in a human-animal conflict to 33 adult male orangutans of similar developmental stage,
84 we found consistent differences between the Batang Toru individual and other extant Ponginae. A
85 second line of evidence provided our analyses of 37 orangutan genomes. Model-based approaches
86 revealed that the deepest split in the evolutionary history of extant orangutans occurred ~3.38 Ma ago
87 between the Batang Toru population and those to the north of Lake Toba, while both currently
88 recognized species separated much later about 674 ka ago. Our combined analyses support a new
89 classification of orangutans into three extant species. The new species, *Pongo tapanuliensis*,
90 encompasses the Batang Toru population, of which fewer than 800 individuals survive.

Results and Discussion

Despite decades of field studies [3] our knowledge of variation among orangutans remains limited as many populations occur in isolated and inaccessible habitats, leaving questions regarding their evolutionary history and taxonomic classification largely unresolved. In particular, Sumatran populations south of Lake Toba had long been overlooked, even though a 1939 review of the species' range mentioned that orangutans had been reported in several forest areas in that region [4]. Based on diverse sources of evidence, we describe a new orangutan species, *Pongo tapanuliensis*, which encompasses a geographically and genetically isolated population found in the Batang Toru area at the southernmost range of extant Sumatran orangutans, south of Lake Toba, Indonesia.

Systematics

Genus *Pongo* Lacépède, 1799

Pongo tapanuliensis sp. nov. Nurcahyo, Meijaard, Nowak, Fredriksson & Groves

Tapanuli Orangutan

Etymology. The species name refers to three North Sumatran districts (North, Central, and South Tapanuli) to which *P. tapanuliensis* is endemic.

Holotype. The complete skeleton of an adult male orangutan that died from wounds sustained by local villagers in November 2013 near Sugi Tonga, Marancar, Tapanuli (Batang Toru) Forest Complex (1°35'54.1"N, 99°16'36.5"E), South Tapanuli District, North Sumatra, Indonesia. Skull and postcranium are lodged in the Museum Zoologicum Bogoriense, Indonesia, accession number MZB39182. High-resolution 3D reconstructions of the skull and mandible are available as supplementary material.

Paratypes. Adult individuals of *P. tapanuliensis* (P2591-M435788 – P2591-M435790) photographed by Tim Laman in the Batang Toru Forest Complex (1°41'9.1"N, 98°59'38.1"E), North Tapanuli District, North Sumatra, Indonesia. Paratypes are available from <http://www.morphobank.org> (Login: 2591 / Password: tapanuliorangutan).

Differential diagnosis. We compared the holotype to a comprehensive comparative data set of 33 adult male orangutans from 10 institutions housing osteological specimens. Unless otherwise stated, all units are in [mm]. Summary statistics for all measurements are listed in Tables S1–3. *Pongo tapanuliensis* differs from all extant orangutans in the breadth of the upper canine (21.5 vs. <20.86); the shallow face depth (6.0 vs. >8.4); the narrower interpterygoid distance (at posterior end of pterygoids 33.8 vs. >43.9; at anterior end of pterygoids, 33.7 vs. >43.0); the shorter tympanic tube (23.9 vs. >28.4, mostly >30); the shorter temporomandibular joint (22.5 vs. >24.7); the narrower maxillary incisor row (28.3 vs. >30.1); the narrower distance across the palate at the first molars (62.7 vs. >65.7); the shorter horizontal

length of the mandibular symphysis (49.3 vs. >53.7); the smaller inferior transverse torus (horizontal length from anterior surface of symphysis 31.8 compared to >36.0); and the width of the ascending ramus of the mandible (55.9 vs. >56.3).

Pongo tapanuliensis differs specifically from *P. abelii* by its deep suborbital fossa, triangular pyriform aperture, and angled facial profile; the longer nuchal surface (70.5 vs. <64.7); the wider rostrum, posterior to the canines (59.9 vs. <59); the narrower orbits (33.8 vs. <34.6); the shorter (29.2 vs. >30.0) and narrower foramen magnum (23.2 vs. >23.3); the narrower bicondylar breadth (120.0 vs. >127.2); the narrower mandibular incisor row (24.4 vs. >28.3); the greater mesio-distal length of the upper canine (19.44 vs. <17.55). The male long call has a higher maximum frequency range of the roar pulse type (>800 Hz vs. <747) with a higher ‘shape’ (>952 Hz/s vs. <934).

Pongo tapanuliensis differs from *P. pygmaeus* by possessing a nearly straight zygomaxillary suture; the lower orbit (orbit height 33.4 vs. >35.3); the male long call has a longer duration (>111 seconds vs. <90) with a greater number of pulses (>52 pulses vs. <45), and is delivered at a greater rate (>0.82 pulses per 20 seconds vs. <0.79).

Pongo tapanuliensis differs specifically from *Pongo ‘pygmaeus’ palaeosumatrensis* in the smaller size of the first upper molar (mesio-distal length 13.65 vs. >14.0, buccolingual breadth 11.37 vs. >12.10, crown area 155.2 mm² vs. >175.45, Figure S1).

Description. Craniometrically, the type skull of *P. tapanuliensis* (Figure 1B) is significantly smaller than any skull of comparable developmental stage of other orangutans; it falls outside of the interquartile ranges of *P. abelii* and *P. pygmaeus* for 24 of 39 cranio-mandibular measurements (Table S1). A principal component analysis (PCA) of 26 cranio-mandibular measurements commonly used in primate taxonomic classification [5, 6] shows consistent differences between *P. tapanuliensis* and the two currently recognized species (Figs. 1C and S2).

The external morphology of *P. tapanuliensis* is more similar to *P. abelii* in its linear body build and more cinnamon pelage than *P. pygmaeus*. The hair texture of *P. tapanuliensis* is frizzier, contrasting in particular with the long, loose body hair of *P. abelii*. *Pongo tapanuliensis* has a prominent moustache and flat flanges covered in downy hair in dominant males, while flanges of older males resemble more those of Bornean males. Females of *P. tapanuliensis* have beards, unlike *P. pygmaeus*.

Distribution. *Pongo tapanuliensis* occurs only in a small number of forest fragments in the districts of Central, North, and South Tapanuli, Indonesia (Figure 1A). The total distribution covers approximately 1,000 km², with an estimated population size of fewer than 800 individuals [7]. The current distribution of *P. tapanuliensis* is almost completely restricted to medium elevation hill and submontane forest (~300–1300 m asl) [7-9]. While densities are highest in primary forest, it does occur at lower densities in mixed agroforest at the edge of primary forest areas [10, 11]. Until relatively recently, *P. tapanuliensis*

was more widespread to the south and west of the current distribution, although evidence for this is largely anecdotal [12, 13].

Other hominoid species and subspecies were previously described using standard univariate and multivariate techniques to quantify morphological character differences. The elevation of bonobos (*P. paniscus*) from a subspecies to a species dates back to Coolidge [14] and was based on summary statistics of primarily morphological data from a single female specimen of *P. paniscus*, five available *P. paniscus* skulls, and comparative data of what is now *P. troglodytes*. Groves and colleagues [5] and Shea et al. [15] supported Coolidge's proposal using larger sample sizes and discriminant function analyses. Shea et al. [15] remarked that the species designation for *P. paniscus*, which was largely based on morphological comparisons, was ultimately strengthened by genetic, ecological, and behavioral data, as we attempted here for *Pongo tapanuliensis*. For the genus *Gorilla*, Stumpf et al. [16] and Groves [17] used cranio-mandibular data from 747 individuals from 19 geographic regions, confirming a classification of the genus into two species (*G. gorilla* and *G. beringei*), as proposed earlier by Groves [1]. Other recent primate species descriptions primarily relied on an inconsistent mix of data on pelage color, ecology, morphology, and/or vocalizations [18-23], with only a few also incorporating genetic analyses [24, 25].

Here, we used an integrative approach by corroborating the morphological analysis, behavioral and ecological data with whole-genome data of 37 orangutans with known provenance, covering the entire range of extant orangutans including areas never sampled before (Figure 2A, Table S4). We applied a model-based approach to statistically evaluate competing demographic models, identify independent evolutionary lineages, and infer levels of gene flow and the timing of genetic isolation between lineages. This enabled us to directly compare complex and realistic models of speciation. We refrained from directly comparing genetic differentiation among the three species in the genus *Pongo* with that of other hominoids, as we deem such comparisons problematic in order to evaluate whether *P. tapanuliensis* constitutes a new species. This is because estimates of genetic differentiation reflect a combination of divergence time, demographic history, and gene flow, and are also influenced by the employed genetic marker system [26, 27].

A PCA (Figure 2B) of genomic diversity highlighted the divergence between individuals from Borneo and Sumatra (PC1), but also separated *P. tapanuliensis* from *P. abelii* (PC2). The same clustering pattern was also found in a model-based analysis of population structure (Figure 2C), and is consistent with an earlier genetic study analyzing a larger number of non-invasively collected samples using microsatellite markers [28]. However, while powerful in detecting extant population structure, population history and speciation cannot be inferred, as they are not suited to distinguish between old divergences with gene flow and cases of recent divergence with isolation [29, 30]. To address this problem and further

investigate the timing of population splits and gene flow, we therefore employed different complementary modeling and phylogenetic approaches.

We applied an Approximate Bayesian Computation (ABC) approach, which allows to infer and compare arbitrarily complex demographic modes based on the comparison of the observed genomic data to extensive population genetic simulations [31]. Our analyses revealed three deep evolutionary lineages in extant orangutans (Figs. 3A and B). Colonization scenarios in which the earliest split within *Pongo* occurred between the lineages leading to *P. abelii* and *P. tapanuliensis* were much better supported than scenarios in which the earliest split was between Bornean and Sumatran species (models 1 vs. models 2, combined posterior probability: 99.91%, Figure 3A). Of the two best scenarios, a model postulating colonization of both northern Sumatra and Borneo from an ancestral population likely situated south of Lake Toba on Sumatra, had the highest support (model 1a vs. model 1b, posterior probability 97.56%, Figure 3A). Our results supported a scenario in which orangutans from mainland Asia first entered Sundaland south of what is now Lake Toba on Sumatra, the most likely entry point based on paleogeographic reconstructions [32]. This ancestral population, of which *P. tapanuliensis* is a direct descendant, then served as a source for the subsequent different colonization events of what is now Borneo, Java and northern Sumatra.

We estimated the split time between populations north and south of Lake Toba at ~3.4 Ma (Figure 3B, Table S5). Under our best-fitting model, we found evidence for post-split gene flow across Lake Toba (~0.3–0.9 migrants per generation, Table S5), which is consistent with highly significant signatures of gene flow between *P. abelii* and *P. tapanuliensis* using D-statistics (CK, BT, WA, *Homo sapiens*: $D = -0.2819$, $p\text{-value} < 0.00001$; WK, BT, LK, *Homo sapiens*: $D = -0.2967$, $p\text{-value} < 0.00001$). Such gene flow resulted in higher autosomal affinity of *P. tapanuliensis* to *P. abelii* compared to *P. pygmaeus* in the PCA (Figure 2B), explaining the smaller amount of variance captured by PC2 (separating *P. tapanuliensis* from all other populations) compared to PC1 (separating *P. pygmaeus* from the Sumatran populations). The parameter estimates from a Bayesian full-likelihood analysis implemented in the software G-PhoCS were in good agreement with those obtained by the ABC analysis, although the split time between populations north and south of Lake Toba was more recent (~2.27 Ma, 95%-HPD: 2.21–2.35, Table S5). The G-PhoCS analysis revealed highly asymmetric gene flow between populations north and south of the Toba caldera, with much lower levels of gene flow into the Batang Toru population from the north than vice versa (Table S5).

The existence of two deep evolutionary lineages among extant Sumatran orangutans was corroborated by phylogenetic analyses based on whole mitochondrial genomes (Figure 4A), in which the deepest split occurred between populations north of Lake Toba and all other orangutans at ~3.97 Ma (95%-HPD: 2.35–5.57). Sumatran orangutans formed a paraphyletic group, with *P. tapanuliensis* being more closely related to the Bornean lineage from which it diverged ~2.41 Ma (1.26–3.42 Ma). In contrast, Bornean

populations formed a monophyletic group with a very recent mitochondrial coalescence at ~160 ka (94–227 ka).

Due to strong female philopatry [33], gene flow in orangutans is almost exclusively male-mediated [34]. Consistent with these pronounced differences in dispersal behavior, phylogenetic analysis of extensive Y-chromosomal sequencing data revealed a comparatively recent coalescence of Y chromosomes of all extant orangutans ~430 ka (Figure 4B). The single available Y-haplotype from *P. tapanuliensis* was nested within the other Sumatran sequences, pointing at the occurrence of male-mediated gene flow across the Toba divide. Thus, in combination with our modeling results, the sex-specific data highlighted the impact of extraordinarily strong male-biased dispersal in the speciation process of orangutans.

Our analyses revealed significant divergence between *P. tapanuliensis* and *P. abelii* (Figs. 3B and 4A), and low levels of male-mediated gene flow (Figs. 3B and 4B), which, however, completely ceased 10–20 ka ago (Figure 3C). Populations north and south of Lake Toba on Sumatra had been in genetic contact for most of the time since their split, but there was a marked reduction in gene flow after ~100 ka (Figure 3C), consistent with habitat destruction caused by the Toba supereruption 73 ka ago [35]. However, *P. tapanuliensis* and *P. abelii* have been on independent evolutionary trajectories at least since the late Pleistocene/early Holocene, as gene flow between these populations has ceased completely 10–20 ka (Figure 3C) and is now impossible because of habitat loss in areas between the species' ranges [7].

Nowadays, most biologists would probably adopt an operational species definition such as: 'a species is a population (or group of populations) with fixed heritable differences from other such populations (or groups of populations)' [36]. With totally allopatric populations, a 'reproductive isolation' criterion, such as is still espoused by adherents of the biological species concept, is not possible [37, 38]. Notwithstanding a long-running debate about the role of gene flow during speciation and genetic interpretations of the species concept [39, 40], genomic studies have found evidence for many instances of recent or ongoing gene flow between taxa which are recognized as distinct and well-established species. This includes examples within each of the other three hominid genera. A recent genomic study using comparable methods to ours revealed extensive gene flow between *Gorilla gorilla* and *G. beringei* until ~20–30 ka [41]. Similar, albeit older and less extensive, admixture occurred between *Pan troglodytes* and *P. paniscus* [42], and was also reported for *Homo sapiens* and *H. neanderthalensis* [43]. *Pongo tapanuliensis* and *P. abelii* appear to be further examples, showing diagnostic phenotypic and other distinctions that had persisted in the past despite gene flow between them.

Due to the challenges involved in collecting suitable specimens for morphological and genomic analyses from critically endangered great apes, our description of *P. tapanuliensis* had to rely on a single skeleton and two individual genomes for our main lines of evidence. When further data will become available, a more detailed picture of the morphological and genomic diversity within this species and of the differences to other *Pongo* species might emerge, which may require further taxonomic revision.

However, is not uncommon to describe species based on a single specimen (*e.g.*, [44-46]), and importantly, there were consistent differences among orangutan populations from multiple independent lines of evidence, warranting the designation of a new species with the limited data at hand.

With a census size of fewer than 800 individuals [7], *P. tapanuliensis* is the least numerous of all great ape species [47]. Its range is located around 200 km from the closest population of *P. abelii* to the north (Figure 2A). A combination of small population size and geographic isolation is of particular high conservation concern, as it may lead to inbreeding depression [48] and threaten population persistence [49]. Highlighting this, we discovered extensive runs of homozygosity in the genomes of both *P. tapanuliensis* individuals (Figure S3), pointing at the occurrence of recent inbreeding.

To ensure long-term survival of *P. tapanuliensis*, conservation measures need to be implemented swiftly. Due to the rugged terrain, external threats have been primarily limited to road construction, illegal clearing of forests, hunting, killings during crop conflict and trade in orangutans [7, 11]. A hydro-electric development has been proposed recently in the area of highest orangutan density, which could impact up to 8% of *P. tapanuliensis*' habitat. This project might lead to further genetic impoverishment and inbreeding, as it would jeopardize chances of maintaining habitat corridors between the western and eastern range (Figure 1A), and smaller nature reserves, all of which maintain small populations of *P. tapanuliensis*.

279 **Author Contributions**

280 Conceived the study and wrote the paper: MPMG, AIN, MK, EM, MGN, CG. Edited the manuscript:
 281 SW, GF, CvS, AS, TMB, DAM, TBS, TD, BG, FC, KSW, EV, PotW, PR, JB, MA, AnN. Carried out
 282 statistical analyses: MPMG, AIN, MGN, AnN, CG, MdM, TD, JA, MDR, AL, MP, JPM, MK, EM, AS,
 283 TMB. Provided samples, and behavioral and ecological data: MGN, MPMG, AnN, AIN, GF, JA, AL,
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References

1. Groves, C.P. (2001). Primate taxonomy, (Washington, D.C. ; London: Smithsonian Institution Press).
2. Xu, X., and Arnason, U. (1996). The mitochondrial DNA molecule of Sumatran orangutan and a molecular proposal for two (Bornean and Sumatran) species of orangutan. *J. Mol. Evol.* 43, 431-437.
3. Wich, S.A., Utami Atmoko, S.S., Mitra Setia, T., and van Schaik, C.P. (2009). Orangutans: geographic variation in behavioral ecology and conservation, (Oxford University Press).
4. Nederlandsch-Indische Vereeniging tot Natuurbescherming (1939). Natuur in Zuid- en Oost-Borneo. Fauna, flora en natuurbescherming in de Zuider- en Ooster-Afdeeling van Borneo. In 3 Jaren Indisch natuur leven. Opstellen over landschappen, dieren en planten, tevens elfde verslag (1936-1938), Nederlandsch-Indische Vereeniging tot Natuurbescherming, ed. (Batavia, Indonesia), pp. 334-411.
5. Groves, C.P., Westwood, C., and Shea, B.T. (1992). Unfinished business - Mahalanobis and a clockwork orang. *J. Hum. Evol.* 22, 327-340.
6. Groves, C.P. (1986). Systematics of the great apes. In Comparative primate biology, Vol.1: Systematics, evolution, and anatomy, D.R. Swindler and J. Erwin, eds. (New York: Alan R. Liss), pp. 187-217.
7. Wich, S.A., Singleton, I., Nowak, M.G., Utami Atmoko, S.S., Nisam, G., Arif, S.M., Putra, R.H., Ardi, R., Fredriksson, G., Usher, G., et al. (2016). Land-cover changes predict steep declines for the Sumatran orangutan (*Pongo abelii*). *Sci. Adv.* 2, e1500789.
8. Laumonier, Y., Uryu, Y., Stüwe, M., Budiman, A., Setiabudi, B., and Hadian, O. (2010). Eco-floristic sectors and deforestation threats in Sumatra: identifying new conservation area network priorities for ecosystem-based land use planning. *Biodivers. Conserv.* 19, 1153-1174.
9. Wich, S.A., Usher, G., Peters, H.H., Khakim, M.F.R., Nowak, M.G., and Fredriksson, G.M. (2014). Preliminary data on the highland Sumatran orangutans (*Pongo abelii*) of Batang Toru. In High Altitude Primates, B.N. Grow, S. Gursky-Doyen and A. Krzton, eds. (New York, NY: Springer New York), pp. 265-283.
10. Meijaard, E. (1997). A survey of some forested areas in South and Central Tapanuli, North Sumatra; new chances for orangutan conservation. (Wageningen: Tropenbos and the Golden Ark).
11. Wich, S.A., Fredriksson, G.M., Usher, G., Peters, H.H., Priatna, D., Basalamah, F., Susanto, W., and Kuhl, H. (2012). Hunting of Sumatran orang-utans and its importance in determining distribution and density. *Biol. Conserv.* 146, 163-169.
12. Kramm, W. (1879). Tochtjes in Tapanoei. *Sumatra-Courant* 20, 1-2.
13. Miller, G.S. (1903). Mammals collected by Dr. W.L. Abbott on the coast and islands of northwest Sumatra. *Proceedings US National Museum, Washington* 26, 437-484.
14. Coolidge, H.J. (1933). *Pan Paniscus*. Pigmy chimpanzee from south of the Congo River. *Am. J. Phys. Anthropol.* 18, 1-59.
15. Shea, B.T., Leigh, S.R., and Groves, C.P. (1993). Multivariate craniometric variation in chimpanzees. In Species, Species Concepts and Primate Evolution, W.H. Kimbel and L.B. Martin, eds. (Boston, MA: Springer US), pp. 265-296.
16. Stumpf, R.M., Polk, J.D., Oates, J.F., Jungers, W.L., Heesy, C.P., Groves, C.P., and Fleagle, J.G. (2002). Patterns of diversity in gorilla cranial morphology. In Gorilla biology: a multidisciplinary perspective, A.B. Taylor and M.L. Goldsmith, eds. (Cambridge: Cambridge University Press), pp. 35-61.
17. Groves, C.P. (2002). A history of gorilla taxonomy. In Gorilla biology: a multidisciplinary perspective, A.B. Taylor and M.L. Goldsmith, eds. (Cambridge: Cambridge University Press), pp. 15-34.
18. Geissmann, T., Lwin, N., Aung, S.S., Aung, T.N., Aung, Z.M., Hla, T.H., Grindley, M., and Momberg, F. (2011). A new species of snub-nosed monkey, genus *Rhinopithecus* Milne-Edwards, 1872 (Primates, Colobinae), from northern Kachin state, northeastern Myanmar. *Am. J. Primatol.* 73, 96-107.

19. Jones, T., Ehardt, C.L., Butynski, T.M., Davenport, T.R., Mpunga, N.E., Machaga, S.J., and De Luca, D.W. (2005). The highland mangabey *Lophocebus kipunji*: a new species of African monkey. *Science* 308, 1161-1164.
20. Li, C., Zhao, C., and Fan, P.F. (2015). White-cheeked macaque (*Macaca leucogenys*): a new macaque species from Medog, southeastern Tibet. *Am. J. Primatol.* 77, 753-766.
21. Munds, R.A., Nekaris, K.A., and Ford, S.M. (2013). Taxonomy of the Bornean slow loris, with new species *Nycticebus kayan* (Primates, Lorisidae). *Am. J. Primatol.* 75, 46-56.
22. Rasoloarison, R.M., Weisrock, D.W., Yoder, A.D., Rakotondravony, D., and Kappeler, P.M. (2013). Two new species of mouse lemurs (Cheirogaleidae: *Microcebus*) from Eastern Madagascar. *Int. J. Primatol.* 34, 455-469.
23. Svensson, M.S., Bersacola, E., Mills, M.S.L., Munds, R.A., Nijman, V., Perkin, A., Masters, J.C., Couette, S., Nekaris, K.A.I., and Bearder, S.K. (2017). A giant among dwarfs: a new species of galago (Primates: Galagidae) from Angola. *Am. J. Phys. Anthropol.* 163, 30-43.
24. Davenport, T.R.B., Stanley, W.T., Sargis, E.J., De Luca, D.W., Mpunga, N.E., Machaga, S.J., and Olson, L.E. (2006). A new genus of African monkey, *Rungwecebus*: Morphology, ecology, and molecular phylogenetics. *Science* 312, 1378-1381.
25. Fan, P.F., He, K., Chen, X., Ortiz, A., Zhang, B., Zhao, C., Li, Y.Q., Zhang, H.B., Kimock, C., Wang, W.Z., et al. (2017). Description of a new species of Hoolock gibbon (Primates: Hylobatidae) based on integrative taxonomy. *Am. J. Primatol.* 79.
26. Jost, L. (2008). G_{st} and its relatives do not measure differentiation. *Mol. Ecol.* 17, 4015-4026.
27. Whitlock, M.C. (2011). G_{st} and D do not replace F_{st}. *Mol. Ecol.* 20, 1083-1091.
28. Nater, A., Arora, N., Greminger, M.P., van Schaik, C.P., Singleton, I., Wich, S.A., Fredriksson, G., Perwitasari-Farajallah, D., Pamungkas, J., and Krützen, M. (2013). Marked population structure and recent migration in the critically endangered Sumatran orangutan (*Pongo abelii*). *J. Hered.* 104, 2-13.
29. Nielsen, R., and Wakeley, J. (2001). Distinguishing migration from isolation: a Markov chain Monte Carlo approach. *Genetics* 158, 885-896.
30. Palsboll, P.J., Berube, M., Aguilar, A., Notarbartolo-Di-Sciara, G., and Nielsen, R. (2004). Discerning between recurrent gene flow and recent divergence under a finite-site mutation model applied to North Atlantic and Mediterranean Sea fin whale (*Balaenoptera physalus*) populations. *Evolution* 58, 670-675.
31. Beaumont, M.A., Zhang, W.Y., and Balding, D.J. (2002). Approximate Bayesian computation in population genetics. *Genetics* 162, 2025-2035.
32. Meijaard, E. (2004). Solving mammalian riddles: a reconstruction of the Tertiary and Quaternary distribution of mammals and their palaeoenvironments in island South-East Asia. (Australian National University), p. 2 v.
33. Arora, N., Van Noordwijk, M.A., Ackermann, C., Willems, E.P., Nater, A., Greminger, M., Nietlisbach, P., Dunkel, L.P., Utami Atmoko, S.S., Pamungkas, J., et al. (2012). Parentage-based pedigree reconstruction reveals female matrilineal clusters and male-biased dispersal in nongregarious Asian great apes, the Bornean orang-utans (*Pongo pygmaeus*). *Molecular ecology* 21, 3352-3362.
34. Nater, A., Nietlisbach, P., Arora, N., van Schaik, C.P., van Noordwijk, M.A., Willems, E.P., Singleton, I., Wich, S.A., Goossens, B., Warren, K.S., et al. (2011). Sex-biased dispersal and volcanic activities shaped phylogeographic patterns of extant orangutans (genus: *Pongo*). *Mol. Biol. Evol.* 28, 2275-2288.
35. Chesner, C.A., Rose, W.I., Deino, A., Drake, R., and Westgate, J.A. (1991). Eruptive history of earth's largest Quaternary caldera (Toba, Indonesia) clarified. *Geology* 19, 200-203.
36. Groves, C.P., and Grubb, P. (2011). Ungulate taxonomy, (Baltimore, Md.: Johns Hopkins University Press).
37. Coyne, J.A., and Orr, H.A. (2004). Speciation, (Sunderland, MA: Sinauer Associates, Inc.).
38. Mayr, E. (1963). Animal species and evolution, (Cambridge,: Belknap Press of Harvard University Press).
39. Arnold, M.L. (2016). Divergence with Genetic Exchange, (Oxford, UK: Oxford University Press).
40. Reznick, D.N., and Ricklefs, R.E. (2009). Darwin's bridge between microevolution and macroevolution. *Nature* 457, 837-842.

- 411 41. Scally, A., Dutheil, J.Y., Hillier, L.W., Jordan, G.E., Goodhead, I., Herrero, J., Hobolth, A.,
412 Lappalainen, T., Mailund, T., Marques-Bonet, T., et al. (2012). Insights into hominid evolution
413 from the gorilla genome sequence. *Nature* 483, 169-175.
- 414 42. de Manuel, M., Kuhlwilm, M., Frandsen, P., Sousa, V.C., Desai, T., Prado-Martinez, J.,
415 Hernandez-Rodriguez, J., Dupanloup, I., Lao, O., Hallast, P., et al. (2016). Chimpanzee genomic
416 diversity reveals ancient admixture with bonobos. *Science* 354, 477.
- 417 43. Kuhlwilm, M., Gronau, I., Hubisz, M.J., de Filippo, C., Prado-Martinez, J., Kircher, M., Fu, Q.,
418 Burbano, H.A., Lalueza-Fox, C., de la Rasilla, M., et al. (2016). Ancient gene flow from early
419 modern humans into Eastern Neanderthals. *Nature* 530, 429-433.
- 420 44. Alba, D.M., Almecija, S., DeMiguel, D., Fortuny, J., de los Rios, M.P., Pina, M., Robles, J.M.,
421 and Moya-Sola, S. (2015). Miocene small-bodied ape from Eurasia sheds light on hominoid
422 evolution. *Science* 350.
- 423 45. Stevens, N.J., Seiffert, E.R., O'Connor, P.M., Roberts, E.M., Schmitz, M.D., Krause, C.,
424 Gorscak, E., Ngasala, S., Hieronymus, T.L., and Temu, J. (2013). Palaeontological evidence for
425 an Oligocene divergence between Old World monkeys and apes. *Nature* 497, 611-614.
- 426 46. Zalmout, I.S., Sanders, W.J., MacLatchy, L.M., Gunnell, G.F., Al-Mufarreah, Y.A., Ali, M.A.,
427 Nasser, A.A.H., Al-Masari, A.M., Al-Sobhi, S.A., Nadhra, A.O., et al. (2010). New Oligocene
428 primate from Saudi Arabia and the divergence of apes and Old World monkeys. *Nature* 466,
429 360-U111.
- 430 47. IUCN (2016). IUCN Red List of Threatened Species. Version 2016.2.
- 431 48. Hedrick, P.W., and Kalinowski, S.T. (2000). Inbreeding depression in conservation biology.
432 *Annu. Rev. Ecol. Syst.* 31, 139-162.
- 433 49. Allendorf, F.W., Luikart, G., and Aitken, S.N. (2013). Conservation and the genetics of
434 populations, 2nd Edition, (Hoboken: John Wiley & Sons).
- 435 50. Locke, D.P., Hillier, L.W., Warren, W.C., Worley, K.C., Nazareth, L.V., Muzny, D.M., Yang,
436 S.-P., Wang, Z., Chinwalla, A.T., Minx, P., et al. (2011). Comparative and demographic analysis
437 of orang-utan genomes. *Nature* 469, 529-533.
- 438 51. Prado-Martinez, J., Sudmant, P.H., Kidd, J.M., Li, H., Kelley, J.L., Lorente-Galdos, B.,
439 Veeramah, K.R., Woerner, A.E., O'Connor, T.D., Santpere, G., et al. (2013). Great ape genetic
440 diversity and population history. *Nature* 499, 471-475.
- 441 52. Arora, N., Nater, A., van Schaik, C.P., Willems, E.P., van Noordwijk, M.A., Goossens, B.,
442 Morf, N., Bastian, M., Knott, C., Morrogh-Bernard, H., et al. (2010). Effects of Pleistocene
443 glaciations and rivers on the population structure of Bornean orangutans (*Pongo pygmaeus*).
444 *Proceedings of the National Academy of Sciences* 107, 21376-21381.
- 445 53. Nater, A., Nietlisbach, P., Arora, N., van Schaik, C.P., van Noordwijk, M.A., Willems, E.P.,
446 Singleton, I., Wich, S.A., Goossens, B., Warren, K.S., et al. (2011). Sex-biased dispersal and
447 volcanic activities shaped phylogeographic patterns of extant orangutans (genus: *Pongo*).
448 *Molecular Biology and Evolution* 28, 2275-2288.
- 449 54. van Noordwijk, M.A., Arora, N., Willems, E.P., Dunkel, L.P., Amda, R.N., Mardianah, N.,
450 Ackermann, C., Krützen, M., and van Schaik, C.P. (2012). Female philopatry and its social
451 benefits among Bornean orangutans. *Behavioral Ecology and Sociobiology* 66, 823-834.
- 452 55. Morrogh-Bernard, H.C., Morf, N.V., Chivers, D.J., and Krützen, M. (2011). Dispersal patterns
453 of orang-utans (*Pongo* spp.) in a Bornean peat-swamp forest. *International Journal of*
454 *Primatology* 32, 362-376.
- 455 56. Nietlisbach, P., Arora, N., Nater, A., Goossens, B., Van Schaik, C.P., and Krützen, M. (2012).
456 Heavily male-biased long-distance dispersal of orang-utans (genus: *Pongo*), as revealed by Y-
457 chromosomal and mitochondrial genetic markers. *Molecular ecology* 21, 3173-3186.
- 458 57. Nater, A., Greminger, M.P., Arora, N., van Schaik, C.P., Goossens, B., Singleton, I., Verschoor,
459 E.J., Warren, K.S., and Krützen, M. (2015). Reconstructing the demographic history of orang-
460 utans using Approximate Bayesian Computation. *Molecular Ecology* 24, 310-327.
- 461 58. Drummond, A.J., Suchard, M.A., Xie, D., and Rambaut, A. (2012). Bayesian phylogenetics
462 with BEAUti and the BEAST 1.7. *Molecular biology and evolution* 29, 1969-1973.
- 463 59. Tamura, K., and Nei, M. (1993). Estimation of the number of nucleotide substitutions in the
464 control region of mitochondrial DNA in humans and chimpanzees. *Molecular Biology and*
465 *Evolution* 10, 512-526.

60. Darriba, D., Taboada, G.L., Doallo, R., and Posada, D. (2012). jModelTest 2: more models, new heuristics and parallel computing. *Nature Methods* 9, 772-772.
61. Röhrer-Ertl, O. (1988). Research history, nomenclature, and taxonomy of the orang-utan. In *Orang-utan Biology*, J. Schwartz, ed. (Oxford, UK: Oxford University Press), pp. 7-18.
62. Shapiro, J.S. (1995). Morphometric variation in the orang utan (*Pongo pygmaeus*), with a comparison of inter- and intraspecific variability in the African apes. Volume PhD Dissertation. (Columbia University).
63. Hooijer, D.A. (1948). Prehistoric teeth of man and of the orang utan from Central Sumatra, with notes on the fossil orang utan from Java and Southern China. *Zool Meded Rijksmus Leiden* 29, 175 - 183.
64. Drawhorn, G.M. (1994). The systematics and Paleodemography of fossil Orangutans (Genus *Pongo*). (University of California).
65. Harrison, T., Jin, C., Zhang, Y., Wang, Y., and Zhu, M. (2014). Fossil *Pongo* from the Early Pleistocene Gigantopithecus fauna of Chongzuo, Guangxi, southern China. *Quaternary International* 354, 59-67.
66. de Vos, J. (1983). The *Pongo* faunas from Java and Sumatra and their significance for biostratigraphical and paleo-ecological interpretations. *Proceedings of the Koninklijke Akademie van Wetenschappen. Series B* 86, 417-425.
67. Bacon, A.-M., Westaway, K., Antoine, P.-O., Düringer, P., Blin, A., Demeter, F., Ponche, J.-L., Zhao, J.-X., Barnes, L.M., Sayavonkhamdy, T., et al. (2015). Late Pleistocene mammalian assemblages of Southeast Asia: New dating, mortality profiles and evolution of the predator-prey relationships in an environmental context. *Palaeogeography, Palaeoclimatology, Palaeoecology* 422, 101-127.
68. Louys, J. (2012). Mammal community structure of Sundanese fossil assemblages from the Late Pleistocene, and a discussion on the ecological effects of the Toba eruption. *Quaternary International* 258, 80-87.
69. Schwartz, J.H., Vu The, L., Nguyen Lan, C., Le Trung, K., and Tattersall, I. (1995). A review of the Pleistocene hominoid fauna of the Socialist Republic of Vietnam (excluding Hylobatidae).
70. Plavcan, J.M. (1994). Comparison of four simple methods for estimating sexual dimorphism in fossils. *Am J Phys Anthropol* 94, 465-476.
71. Greminger, M.P., Stolting, K., Nater, A., Goossens, B., Arora, N., Bruggmann, R., Patrignani, A., Nussberger, B., Sharma, R., Kraus, R.H., et al. (2014). Generation of SNP datasets for orangutan population genomics using improved reduced-representation sequencing and direct comparisons of SNP calling algorithms. *BMC genomics* 15, 16.
72. Andrews, S. (2012). FastQC. A quality control tool for high throughput sequence data.
73. Li, H., and Durbin, R. (2009). Fast and accurate short read alignment with Burrows-Wheeler transform. *Bioinformatics* 25, 1754-1760.
74. McKenna, A., Hanna, M., Banks, E., Sivachenko, A., Cibulskis, K., Kernysky, A., Garimella, K., Altshuler, D., Gabriel, S., Daly, M., et al. (2010). The Genome Analysis Toolkit: A MapReduce framework for analyzing next-generation DNA sequencing data. *Genome Research* 20, 1297-1303.
75. DePristo, M.A., Banks, E., Poplin, R., Garimella, K.V., Maguire, J.R., Hartl, C., Philippakis, A.A., del Angel, G., Rivas, M.A., Hanna, M., et al. (2011). A framework for variation discovery and genotyping using next-generation DNA sequencing data. *Nat Genet* 43, 491-498.
76. Derrien, T., Estellé, J., Marco Sola, S., Knowles, D.G., Raineri, E., Guigó, R., and Ribeca, P. (2012). Fast Computation and Applications of Genome Mappability. *PLoS ONE* 7, e30377.
77. Auton, A., and McVean, G. (2007). Recombination rate estimation in the presence of hotspots. *Genome Research* 17, 1219-1227.
78. Auton, A., Fledel-Alon, A., Pfeifer, S., Venn, O., Segurel, L., Street, T., Leffler, E.M., Bowden, R., Aneas, I., Broxholme, J., et al. (2012). A fine-scale chimpanzee genetic map from population sequencing. *Science* 336, 193-198.
79. Delaneau, O., Marchini, J., and Zagury, J.F. (2012). A linear complexity phasing method for thousands of genomes. *Nat Methods* 9, 179-181.
80. Delaneau, O., Howie, B., Cox, A.J., Zagury, J.F., and Marchini, J. (2013). Haplotype estimation using sequencing reads. *American Journal of Human Genetics* 93, 687-696.

- 522 81. McQuillan, R., Leutenegger, A.L., Abdel-Rahman, R., Franklin, C.S., Pericic, M., Barac-Lauc,
523 L., Smolej-Narancic, N., Janicijevic, B., Polasek, O., Tenesa, A., et al. (2008). Runs of
524 homozygosity in European populations. *American Journal of Human Genetics* 83, 359-372.
- 525 82. Pemberton, Trevor J., Absher, D., Feldman, Marcus W., Myers, Richard M., Rosenberg,
526 Noah A., and Li, Jun Z. (2012). Genomic Patterns of Homozygosity in Worldwide Human
527 Populations. *The American Journal of Human Genetics* 91, 275-292.
- 528 83. Hall, T.A. (1999). BioEdit: a user-friendly biological sequence alignment editor and analysis
529 program for Windows 95/98/NT. In *Nucleic acids symposium series*, Volume 41. pp. 95-98.
- 530 84. Roos, C., Zinner, D., Kubatko, L., Schwarz, C., Yang, M., Meyer, D., Nash, S., Xing, J., Batzer,
531 M., Brameier, M., et al. (2011). Nuclear versus mitochondrial DNA: evidence for hybridization
532 in colobine monkeys. *BMC Evolutionary Biology* 11, 77.
- 533 85. Thalmann, O., Serre, D., Hofreiter, M., Lukas, D., Eriksson, J., and Vigilant, L. (2005). Nuclear
534 insertions help and hinder inference of the evolutionary history of gorilla mtDNA. *Molecular*
535 *Ecology* 14, 179-188.
- 536 86. Steiper, M.E., and Young, N.M. (2006). Primate molecular divergence dates. *Molecular*
537 *phylogenetics and evolution* 41, 384-394.
- 538 87. Bellott, D.W., Hughes, J.F., Skaletsky, H., Brown, L.G., Pyntikova, T., Cho, T.-J., Koutseva,
539 N., Zaghoul, S., Graves, T., and Rock, S. (2014). Mammalian Y chromosomes retain widely
540 expressed dosage-sensitive regulators. *Nature* 508, 494-499.
- 541 88. Soh, Y.S., Alföldi, J., Pyntikova, T., Brown, L.G., Graves, T., Minx, P.J., Fulton, R.S.,
542 Kremitzki, C., Koutseva, N., and Mueller, J.L. (2014). Sequencing the mouse Y chromosome
543 reveals convergent gene acquisition and amplification on both sex chromosomes. *Cell* 159, 800-
544 813.
- 545 89. Hughes, J.F., Skaletsky, H., Pyntikova, T., Graves, T.A., van Daalen, S.K., Minx, P.J., Fulton,
546 R.S., McGrath, S.D., Locke, D.P., and Friedman, C. (2010). Chimpanzee and human Y
547 chromosomes are remarkably divergent in structure and gene content. *Nature* 463, 536-539.
- 548 90. Wei, W., Ayub, Q., Chen, Y., McCarthy, S., Hou, Y., Carbone, I., Xue, Y., and Tyler-Smith, C.
549 (2013). A calibrated human Y-chromosomal phylogeny based on resequencing. *Genome*
550 *research* 23, 388-395.
- 551 91. Li, H., Handsaker, B., Wysoker, A., Fennell, T., Ruan, J., Homer, N., Marth, G., Abecasis, G.,
552 Durbin, R., and Subgroup, G.P.D.P. (2009). The Sequence Alignment/Map format and
553 SAMtools. *Bioinformatics* 25, 2078-2079.
- 554 92. Danecek, P., Auton, A., Abecasis, G., Albers, C.A., Banks, E., DePristo, M.A., Handsaker, R.E.,
555 Lunter, G., Marth, G.T., Sherry, S.T., et al. (2011). The variant call format and VCFtools.
556 *Bioinformatics* 27, 2156-2158.
- 557 93. Tavaré, S. (1986). Some probabilistic and statistical problems in the analysis of DNA sequences.
558 In *Lectures on Mathematics in the Life Sciences*, Volume 17. pp. 57-86.
- 559 94. Posada, D. (2003). Using MODELTEST and PAUP* to Select a Model of Nucleotide
560 Substitution. In *Current Protocols in Bioinformatics*. (John Wiley & Sons, Inc.).
- 561 95. Drummond, A.J., Ho, S.Y., Phillips, M.J., and Rambaut, A. (2006). Relaxed phylogenetics and
562 dating with confidence. *PLoS biology* 4, e88.
- 563 96. Yang, Z., and Rannala, B. (2006). Bayesian estimation of species divergence times under a
564 molecular clock using multiple fossil calibrations with soft bounds. *Molecular biology and*
565 *evolution* 23, 212-226.
- 566 97. Brunet, M., Guy, F., Pilbeam, D., Mackaye, H.T., Likius, A., Ahounta, D., Beauvilain, A.,
567 Blondel, C., Bocherens, H., and Boisserie, J.-R. (2002). A new hominid from the Upper
568 Miocene of Chad, Central Africa. *Nature* 418, 145-151.
- 569 98. Vignaud, P., Düringer, P., Mackaye, H.T., Likius, A., Blondel, C., Boisserie, J.-R., De Bonis,
570 L., Eisenmann, V., Etienne, M.-E., and Geraads, D. (2002). Geology and palaeontology of the
571 Upper Miocene Toros-Menalla hominid locality, Chad. *Nature* 418, 152-155.
- 572 99. Raaum, R.L., Sterner, K.N., Noviello, C.M., Stewart, C.-B., and Disotell, T.R. (2005).
573 Catarrhine primate divergence dates estimated from complete mitochondrial genomes:
574 concordance with fossil and nuclear DNA evidence. *J Hum Evol* 48, 237-257.
- 575 100. Rambaut, A., Suchard, M.A., Xie, D., and Drummond, A.J. (2014). Tracer v1.6.
- 576 101. Rambaut, A. (2012). FigTree version 1.4.

- 577 102. Tamura, K., Stecher, G., Peterson, D., Filipski, A., and Kumar, S. (2013). MEGA6: Molecular
578 Evolutionary Genetics Analysis Version 6.0. *Molecular biology and evolution* 30, 2725-2729.
- 579 103. Scally, A., and Durbin, R. (2012). Revising the human mutation rate: implications for
580 understanding human evolution. *Nature Reviews Genetics* 13, 745-753.
- 581 104. Ségurel, L., Wyman, M.J., and Przeworski, M. (2014). Determinants of Mutation Rate Variation
582 in the Human Germline. *Annual Review of Genomics and Human Genetics* 15, 47-70.
- 583 105. Venn, O., Turner, I., Mathieson, I., de Groot, N., Bontrop, R., and McVean, G. (2014). Strong
584 male bias drives germline mutation in chimpanzees. *Science* 344, 1272-1275.
- 585 106. Lipson, M., Loh, P.-R., Sankararaman, S., Patterson, N., Berger, B., and Reich, D. (2015).
586 Calibrating the human mutation rate via ancestral recombination density in diploid genomes.
587 *PLoS Genet* 11, e1005550.
- 588 107. Carbone, L., Alan Harris, R., Gnerre, S., Veeramah, K.R., Lorente-Galdos, B., Huddleston, J.,
589 Meyer, T.J., Herrero, J., Roos, C., Aken, B., et al. (2014). Gibbon genome and the fast karyotype
590 evolution of small apes. *Nature* 513, 195-201.
- 591 108. Wich, S., De Vries, H., Ancrenaz, M., Perkins, L., Shumaker, R., Suzuki, A., and Van Schaik,
592 C. (2009). Orangutan life history variation. In *Orangutans - Geographic Variation in Behavioral*
593 *Ecology and Conservation* S.A. Wich, S.S. Utami Atmoko, T. Mitra Setia and C.P. van Schaik,
594 eds. (Oxford University Press), pp. 65-75.
- 595 109. Team, R.D.C. (2010). R: a language and environment for statistical computing. (Vienna,
596 Austria: R Foundation for Statistical Computing).
- 597 110. Alexander, D.H., Novembre, J., and Lange, K. (2009). Fast model-based estimation of ancestry
598 in unrelated individuals. *Genome Research* 19, 1655-1664.
- 599 111. Purcell, S., Neale, B., Todd-Brown, K., Thomas, L., Ferreira, M.A., Bender, D., Maller, J.,
600 Sklar, P., de Bakker, P.I., Daly, M.J., et al. (2007). PLINK: a tool set for whole-genome
601 association and population-based linkage analyses. *Am J Hum Genet* 81, 559-575.
- 602 112. Schiffels, S., and Durbin, R. (2014). Inferring human population size and separation history
603 from multiple genome sequences. *Nat. Genet.* 46, 919-925.
- 604 113. Robinson, J.D., Bunnefeld, L., Hearn, J., Stone, G.N., and Hickerson, M.J. (2014). ABC
605 inference of multi-population divergence with admixture from unphased population genomic
606 data. *Mol. Ecol.* 23, 4458-4471.
- 607 114. Excoffier, L., Smouse, P.E., and Quattro, J.M. (1992). Analysis of molecular variance inferred
608 from metric distances among DNA haplotypes - application to human mitochondrial DNA
609 restriction data. *Genetics* 131, 479-491.
- 610 115. Hudson, R.R. (2002). Generating samples under a Wright-Fisher neutral model of genetic
611 variation. *Bioinformatics* 18, 337-338.
- 612 116. Le Cao, K.A., Gonzalez, I., and Dejean, S. (2009). integrOmics: an R package to unravel
613 relationships between two omics datasets. *Bioinformatics* 25, 2855-2856.
- 614 117. Csillery, K., Francois, O., and Blum, M.G.B. (2012). abc: an R package for approximate
615 Bayesian computation (ABC). *Methods Ecol. Evol.* 3, 475-479.
- 616 118. Mevik, B.H., and Wehrens, R. (2007). The pls package: principal component and partial least
617 squares regression in R. *J. Stat. Softw.* 18.
- 618 119. Wegmann, D., Leuenberger, C., and Excoffier, L. (2009). Efficient Approximate Bayesian
619 computation coupled with Markov chain Monte Carlo without likelihood. *Genetics* 182, 1207-
620 1218.
- 621 120. Leuenberger, C., and Wegmann, D. (2010). Bayesian computation and model selection without
622 likelihoods. *Genetics* 184, 243-252.
- 623 121. Wegmann, D., Leuenberger, C., Neuenschwander, S., and Excoffier, L. (2010). ABCtoolbox: a
624 versatile toolkit for approximate Bayesian computations. *BMC Bioinformatics* 11.
- 625 122. Cook, S.R., Gelman, A., and Rubin, D.B. (2006). Validation of software for Bayesian models
626 using posterior quantiles. *J. Comput. Graph. Stat.* 15, 675-692.
- 627 123. Rice, W.R. (1989). Analyzing tables of statistical tests. *Evolution* 43, 223-225.
- 628 124. Gronau, I., Hubisz, M.J., Gulko, B., Danko, C.G., and Siepel, A. (2011). Bayesian inference of
629 ancient human demography from individual genome sequences. *Nat. Genet.* 43, 1031-1034.
- 630 125. Baele, G., Lemey, P., Bedford, T., Rambaut, A., Suchard, M.A., and Alekseyenko, A.V. (2012).
631 Improving the accuracy of demographic and molecular clock model comparison while
632 accommodating phylogenetic uncertainty. *Mol Biol Evol* 29, 2157-2167.

126. Raftery, A.E., Newton, M.A., Satagopan, J.M., and Krivitsky, P.N. (2007). Estimating the integrated likelihood via posterior simulation using the harmonic mean identity. In Bayesian Statistics, J.M. Bernardo, M.J. Bayarri and J.O. Berger, eds. (Oxford: Oxford University Press), pp. 1-45.
127. Röhrer-Ertl, O. (1984). Orang-utan Studien, (Neuried, Germany: Hieronymus Verlag).
128. Röhrer-Ertl, O. (1988). Cranial growth. In Orang-utan Biology, J. Schwartz, ed. (Oxford, UK: Oxford University Press), pp. 201-224.
129. Courtenay, J., Groves, C., and Andrews, P. (1988). Inter- or intra-island variation? An assessment of the differences between Bornean and Sumatran orang-utans. In Orang-utan biology, H. Schwartz, ed. (Oxford, England: Oxford University Press), pp. 19-29.
130. Uchida, A. (1998). Variation in tooth morphology of *Pongo pygmaeus*. *J Hum Evol* 34, 71-79.
131. Taylor, A.B. (2006). Feeding behavior, diet, and the functional consequences of jaw form in orangutans, with implications for the evolution of Pongo. *J Hum Evol* 50, 377-393.
132. Taylor, A.B. (2009). The functional significance of variation in jaw form in orangutans. In Orangutans: geographic variation in behavioral ecology and conservation, S.A. Wich, S.U. Atmoko, T.M. Setia and C.P. van Schaik, eds. (Oxford, UK.: Oxford University Press), pp. 15-31.
133. Tukey, J.W. (1977). Exploratory data analysis, (London, UK: Addison-Wesley Publishing Company).
134. Tabachnick, B.G., and Fidell, L.S. (2013). Using multivariate statistics, 6th ed, (New York, USA: Pearson).
135. R Core Development Team (2016). R: A language and environment for statistical computing. R Foundation for Statistical Computing. <http://www.R-project.org/>. (Vienna, Austria).
136. Kaiser, H.F. (1960). The application of electronic computers to factor analysis. *Education and Psychological Measurement* 20, 141-151.
137. Revelle, W. (2016). Psych: procedures for personality and psychological research. <http://CRAN.R-project.org/package=psych> Version =1.6.4, (Evanston, Illinois, USA: Northwestern University).
138. Davila-Ross, M. (2004). The long calls of wild male orangutans: A phylogenetic approach. Volume PhD. (Hannover, Germany: Institut für Zoologie, Tierärztliche Hochschule Hannover).
139. Davila-Ross, M., and Geissmann, T. (2007). Call diversity of wild male orangutans: a phylogenetic approach. *Am. J. Primatol.* 69, 305-324.
140. Lameira, A.R., and Wich, S.A. (2008). Orangutan Long Call Degradation and Individuality Over Distance: A Playback Approach. *Int. J. Primatol.* 29, 615-625.
141. Delgado, R.A., Lameira, A.R., Davila Ross, M., Husson, S.J., Morrogh-Bernard, H.C., and Wich, S.A. (2009). Geographical variation in orangutan long calls. In Orangutans: Geographic variation in behavioral ecology and conservation, S.A. Wich, S.S. Utami Atmoko, T. Mitra Setia and C.P. van Schaik, eds. (Oxford, UK: Oxford University Press), pp. 215-224.
142. Darul Sukma, W.P., Dai, J., Hidayat, A., Yayat, A.H., Sumulyadi, H.Y., Hendra, S., Buurman, P., and Balsem, T. (1990). Explanatory booklet of the land unit and soil map of the Sidikalang sheet (618), Sumatra. (Bogor, Indonesia: Centre for Soil and Agroclimate Research).
143. Darul Sukma, W.P., Suratman, Hidayat, J.A., and Budhi, P.G. (1990). Explanatory booklet of the land unit and soil map of the Tapaktuan sheet (519), Sumatra, (Bogor, Indonesia: Centre for Soil and Agroclimate Research).
144. Darul Sukma, W.P., Suratman, Hidayat, J.A., and Budi, P.G. (1990). Explanatory booklet of the land unit and soil map of the Lho'Kruet sheet (420), Sumatra, (Bogor, Indonesia: Centre for Soil and Agroclimate Research).
145. Darul Sukma, W.P., Verhagen, V., Dai, J., Buurman, P., Balsem, T., Suratman, and Vejre, H. (1990). Explanatory booklet of the land unit and soil map of the Takengon sheet (520), Sumatra, (Bogor, Indonesia: Centre for Soil and Agroclimate Research).
146. Hidayat, A., Verhagen, A., Darul Sukma, W.P., Buurman, P., Balsem, T., Suratman, and Vejre, H. (1990). Explanatory booklet of the land unit and soil map of the Lhokseumawe (521) and Simpangulim (621) sheets, Sumatra, (Centre for Soil and Agroclimate Research).
147. Hikmatullah, Wahyunto, Chendy, T.F., Dai, J., and Hidayat, A. (1990). Explanatory booklet of the land unit and soil map of the Langsa (620) sheet, Sumatra, (Bogor, Indonesia: Centre for Soil and Agroclimate Research).

148. Subardja, D., Djuanda, K., Hadian, Y., Samdan, C.D., Mulyadi, Y., Supriatna, W., and Dai, J. (1990). Explanatory booklet of the land unit and soil map of the Sibolga (617) and Padangsidempuan (717) sheets, Sumatra, (Bogor, Indonesia: Centre for Soil and Agroclimate Research).
149. Wahyunto, Puksi, D.S., Rochman, A., Wahdini, W., Paidi, Dai, J., Hidayat, A., Buurman, P., and Balsem, T. (1990). Explanatory booklet of the land unit and soil map of the Medan (619) sheet, Sumatra, (Bogor, Indonesia: Centre for Soil and Agroclimate Research).
150. Hall, R., van Hattum, M.W.A., and Spakman, W. (2008). Impact of India–Asia collision on SE Asia: The record in Borneo. *Tectonophysics* 451, 366-389.
151. Hijmans, R.J., Cameron, S.E., Parra, J.L., Jones, P.G., and Jarvis, A. (2005). Very high resolution interpolated climate surfaces for global land areas. *International Journal of Climatology* 25, 1965-1978.
152. Wich, S.A., Singleton, I., Nowak, M.G., Utami Atmoko, S.S., Nisam, G., Arif, S.M., Putra, R.H., Ardi, R., Fredriksson, G., Usher, G., et al. (2016). Land-cover changes predict steep declines for the Sumatran orangutan (*Pongo abelii*). *Science Advances* 2, e1500789.
153. Wich, S.A., Atmoko, S.U., Setia, T.M., and van Schaik, C. (2009). Orangutans. Geographic variation in behavioral ecology and conservation, (Oxford, UK: Oxford University Press).
154. Hall, T.A. (1999). BioEdit: a user-friendly biological sequence alignment editor and analysis program for Windows 95/98/NT. *Nucleic acids symposium series* 41, 95-98.
155. Li, H., Handsaker, B., Wysoker, A., Fennell, T., Ruan, J., Homer, N., Marth, G., Abecasis, G., and Durbin, R. (2009). The sequence alignment/map format and SAMtools. *Bioinformatics* 25, 2078-2079.
156. Danecek, P., Auton, A., Abecasis, G., Albers, C.A., Banks, E., DePristo, M.A., Handsaker, R.E., Lunter, G., Marth, G.T., Sherry, S.T., et al. (2011). The variant call format and VCFtools. *Bioinformatics* 27, 2156-2158.
157. Patterson, N., Moorjani, P., Luo, Y., Mallick, S., Rohland, N., Zhan, Y., Genschoreck, T., Webster, T., and Reich, D. (2012). Ancient admixture in human history. *Genetics* 192, 1065-1093.
158. Venables, W.N., and Ripley, B.D. (2002). *Modern applied statistics with S*, 4th edition, (New York, USA: Springer).

Figure 1. Morphological evidence supporting a new orangutan species. A) Current distribution of *Pongo tapanuliensis* on Sumatra. The holotype locality is marked with a red star. The area shown in the map is indicated in Figure 2A. B) Holotype skull and mandible of *P. tapanuliensis* from a recently deceased individual from Batang Toru. See also Figure S1, Tables S1 and S2. C) Violin plots of the first seven principal components of 26 cranio-mandibular morphological variables of 8 north Sumatran *P. abelii* and 19 Bornean *P. pygmaeus* individuals of similar developmental state as the holotype skull (black horizontal lines). See also Figure S2.

Figure 2. Distribution, genomic diversity, and population structure of the genus *Pongo*. A) Sampling areas across the current distribution of orangutans. The contour indicates the extent of the exposed Sunda Shelf during the last glacial maximum. The black rectangle delimits the area shown in Figure 1A. n = numbers of sequenced individuals. See also Table S4. B) Principal component analysis of genomic diversity in *Pongo*. Axis labels show the percentages of the total variance explained by the first two principal components. Colored bars in the insert represent the distribution of nucleotide diversity in genome-wide 1-Mb windows across sampling areas. C) Bayesian clustering analysis of population structure using the program ADMIXTURE. Each vertical bar depicts an individual, with colors representing the inferred ancestry proportions with different assumed numbers of genetic clusters (K, horizontal sections).

Figure 3. Demographic history and gene flow in *Pongo*. A) Model selection by Approximate Bayesian Computation (ABC) of plausible colonization histories of orangutans on Sundaland. The ABC analyses are based on the comparison of ~3,000 non-coding 2-kb loci randomly distributed across the genome with corresponding data simulated under the different demographic models. The numbers in the black boxes indicate the model's posterior probability. NT = Sumatran populations north of Lake Toba, ST = the Sumatran population of Batang Toru south of Lake Toba, BO = Bornean populations. B) ABC parameter estimates based on the full demographic model with colonization pattern inferred in panel A. Numbers in grey rectangles represent point estimates of effective population size (N_e). Arrows indicate gene flow among populations, numbers above the arrows represent point estimates of numbers of migrants per generation. See also Table S5. C) Relative cross-coalescent rate (RCCR) analysis for between-species pairs of phased high-coverage genomes. A RCCR close to 1 indicates extensive gene flow between species, while a ratio close to 0 indicates genetic isolation between species pairs. The x-axis shows time scaled in years, assuming a generation time of 25 years and an autosomal mutation rate of 1.5×10^{-8} per site per generation. See also Figure S3.

Figure 4. Sex-specific evolutionary history of orangutans. Bayesian phylogenetic trees for (A) mitochondrial genomes and (B) Y chromosomes. The mitochondrial tree is rooted with a human and a central chimpanzee sequence, the Y chromosome tree with a human sequence (not shown). ** Posterior probability = 1.00. C) Genotype-sharing matrix for mitogenomes (above the diagonal) and Y

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754 chromosomes (below the diagonal) for all analyzed male orangutans. A value of 1 indicates that two
755 males have identical genotypes at all polymorphic sites; a value of 0 means that they have different
756 genotypes at all variable positions.

CONTACT FOR RESOURCE SHARING

Further information and requests for resources and reagents should be directed to and will be fulfilled by the Lead Contact, Michael Krützen (michael.krutzen@aim.uzh.ch).

EXPERIMENTAL MODEL AND SUBJECT DETAILS

Sample collection and population assignment for genomic analysis

Our sample set comprised genomes from 37 orangutans, representing the entire geographic range of extant orangutans (Figure 2A). We obtained whole-genome sequencing data for the study individuals from three different sources (Table S4): (i) genomes of 17 orangutans were sequenced for this study. Data for 20 individuals were obtained from (ii) Locke *et al.* [50] (n=10) and (iii) Prado-Martinez *et al.* [51] (n=10). All individuals were wild-born, except for five orangutans which were first-generation offspring of wild-born parents of the same species (Table S4).

Population provenance of the previously sequenced orangutans [50, 51] was largely unknown. We identified their most likely natal area based on mtDNA haplotype clustering in a phylogenetic tree together with samples of known geographic provenance. Because of extreme female philopatry in orangutans, mtDNA haplotypes are reliable indicators for the population of origin [33, 52-56]. Using three concatenated mtDNA genes (16S ribosomal DNA, Cytochrome b, and NADH-ubiquinone oxidoreductase chain 3), we constructed a Bayesian tree, including 127 non-invasively sampled wild orangutans from 15 geographic regions representing all known extant orangutan populations [53, 57]. Gene sequences of our study individuals were extracted from their complete mitochondrial genome sequences. The phylogenetic tree was built with BEAST v1.8.0. [58], as described in Nater *et al.* [53], applying a TN93+I substitution model [59] as determined by jModelTest v2.1.4. [60].

Using the mitochondrial tree, we assigned all previously sequenced orangutans [50, 51] to their most likely population of origin. Our sample assignment revealed incomplete geographic representation of the genus *Pongo* in previous studies. To achieve a more complete representation of extant orangutans, we sequenced genomes of 17 wild-born orangutans mainly from areas with little or no previous sample coverage. Detailed provenance information for these individuals is provided in Table S4.

Samples for morphological analysis

We conducted comparative morphological analyses of 34 adult male orangutans from 10 institutions housing osteological specimens. A single adult male skeleton from the Batang Toru population was available for study, having died from injuries sustained in an orangutan-human conflict situation in November 2013. To account for potential morphological differences related to developmental stage [61, 62], our analyses included only males at a similar developmental stage as the Batang Toru specimen,

i.e., having a sagittal crest of <10 mm in height. In addition to the single available Batang Toru male, our extant sample comprises specimens from the two currently recognized species, the north Sumatran *Pongo abelii* (n=8) and the Bornean *P. pygmaeus* (n=25).

We also evaluated the relationship of the dental material between the Batang Toru specimen and those of the Late Pleistocene fossil material found within the Djamboe, Lida Ajer, and Sibrambang caves near Padang, Sumatra, all of which has been previously described by Hooijer [63]. Some scholars have suggested that the fossil material may represent multiple species [64, 65]. However, Hooijer had more than adequately shown that the variation in dental morphology observed within the three cave assemblages can easily be accommodated within a single species [63]. As only teeth were present in the described cave material, many of which also have gnaw marks, taphonomic processes (*e.g.*, porcupines as accumulating agents) are thought to have largely shaped the cave material [66, 67] and thus may account for the appearance of size differences among the cave samples [64, 65]. Furthermore, the similarities in the reconstructed age of the cave material (~128-118 ka or ~80-60 ka [66-68]), and the fact that the presence of more than one large-bodied ape species is an uncommon feature in both fossil and extant Southeast Asian faunal assemblages [69], makes it highly unlikely that multiple large-bodied ape species co-existed within the area at a given time. For purposes of discussion here, we collectively refer to the Padang fossil material as *P. p. palaeosumatrensis*, as described by Hooijer [63].

As the comparative fossil sample likely comprises various age-sex classes [63], we divided the fossil sample into two portions above and below the mean for each respective tooth utilized in this study. We considered samples above the mean to represent larger individuals, which we attribute to “males”, and the ones below to being smaller individuals, which we attribute to “females” [70]. We only used the “male” samples in comparison to our extant male comparative orangutan sample.

METHOD DETAILS

Whole-genome sequencing

To obtain sufficient amounts of DNA, we collected blood samples from confiscated orangutans at rehabilitation centres, including the Sumatran Orangutan Conservation Program (SOCP) in Medan, BOS Wanariset Orangutan Reintroduction Project in East Kalimantan, Semongok Wildlife Rehabilitation Centre in Sarawak, and Sepilok Orangutan Rehabilitation Centre in Sabah. We took whole blood samples during routine veterinary examinations and stored in EDTA blood collection tubes at -20°C. The collection and transport of samples were conducted in strict accordance with Indonesian, Malaysian and international regulations. Samples were transferred to Zurich under the Convention on International Trade of Endangered Species in Fauna and Flora (CITES) permit numbers 4872/2010 (Sabah), and 06968/IV/SATS-LN/2005 (Indonesia).

We extracted genomic DNA using the Gentra Puregene Blood Kit (Qiagen) but modified the protocol for clotted blood as described in Greminger *et al.* [71]. We sequenced individuals on two to three lanes on an Illumina HiSeq 2000 in paired end (2 x 101 bp) mode. Sample PP_5062 was sequenced at the Functional Genomics Center in Zurich (Switzerland), the other individuals at the Centre Nacional d'Anàlisi Genòmica in Barcelona (Spain), as the individuals of Prado-Martinez *et al.* [51]. On average, we generated $\sim 1.1 \times 10^9$ raw Illumina reads per individual.

Read mapping

We followed identical bioinformatical procedures for all 37 study individuals, using the same software versions. We quality-checked raw Illumina sequencing reads with FastQC v0.10.1. [72] and mapped to the orangutan reference genome *ponAbe2* [50] using the Burrows-Wheeler Aligner (BWA-MEM) v0.7.5 [73] in paired-end mode with default read alignment penalty scores. We used Picard v1.101 (<http://picard.sourceforge.net/>) to add read groups, convert sequence alignment/map (SAM) files to binary alignment/map (BAM) files, merge BAM files for each individual, and to mark optical and PCR duplicates. We filtered out duplicated reads, bad read mates, reads with mapping quality zero, and reads that mapped ambiguously.

We performed local realignment around indels and empirical base quality score recalibration (BQSR) with the Genome Analysis Toolkit (GATK) v3.2.2. [74, 75]. The BQSR process empirically calculates more accurate base quality scores (*i.e.*, Phred-scaled probability of error) than those emitted by the sequencing machines through analysing the covariation among several characteristics of a base (*e.g.*, position within the read, sequencing cycle, previous base, etc.) and its status of matching the reference sequence or not. To account for true sequence variation in the data set, the model requires a database of known polymorphic sites ('known sites') which are skipped over in the recalibration algorithm. Since no suitable set of 'known sites' was available for the complete genus *Pongo*, we preliminary identified

confident SNPs from our data. For this, we performed an initial round of SNP calling on unrecalibrated BAM files with the *UnifiedGenotyper* of the GATK. Single nucleotide polymorphisms were called separately for Bornean and Sumatran orangutans in multi-sample mode (*i.e.*, joint analysis of all individuals per island), creating two variant call (VCF) files. In addition, we produced a third VCF file jointly analysing all study individuals in order to capture genus-wide low frequency alleles. We applied the following hard quality filter criteria on all three VCF files: $QUAL < 50.0 \parallel QD < 2.0 \parallel FS > 60.0 \parallel MQ < 40.0 \parallel HaplotypeScore > 13.0 \parallel MappingQualityRankSum < -12.5 \parallel ReadPosRankSum < -8.0$. Additionally, we calculated the mean and standard deviation of sequencing depth over all samples and filtered all sites with a site-wise coverage more than five standard deviations above the mean. We merged the three hard filtered VCF files and took SNPs as ‘known sites’ for BQSR with the GATK. The walkers CountReads and DepthOfCoverage of the GATK were used to obtain various mapping statistics for unfiltered and filtered BAM files.

Mean effective sequencing depth, estimated from filtered BAM files, varied among individuals ranging from 4.8–12.2x [50] to 13.7–31.1x (this study) [51], with an average depth of 18.4x over all individuals (Tables S4). For the previously sequenced genomes [50, 51], estimated sequence depths were 25–40% lower as the values reported in the two source studies. This difference is explained by the way sequence depth was calculated. Here, we estimated sequence depth on the filtered BAM files where duplicated reads, bad read mates, reads with mapping quality zero, and reads which mapped ambiguously had already been removed. Thus, our sequence coverage estimates correspond to the effective read-depths which are available for SNP discovery and genotyping.

SNP and genotype calling

We produced high quality genotypes for all individuals for each position in the genome, applying the same filtering criteria for SNP and non-polymorphic positions. We identified SNPs and called genotypes in a three-step approach. First, we identified a set of candidate (raw) SNPs among all study individuals. Second, we performed variant quality score recalibration (VQSR) on the candidate SNPs to identify high-confidence SNPs. Third, we called genotypes of all study individuals at these high-confidence SNP positions.

Step 1: We used the *HaplotypeCaller* of the GATK in genomic Variant Call Format (gVCF) mode to obtain for each individual in the dataset genotype likelihoods at any site in the reference genome. *HaplotypeCaller* performs local realignment of reads around potential variant sites and is therefore expected to considerably improve SNP calling in difficult-to-align regions of the genome. We then genotyped the resulting gVCF files together on a per-island level, as well as combined for all individuals, using the *Genotype GVCFs* tool of the GATK to obtain three VCF files with candidate SNPs for *P. abelii*, *P. pygmaeus*, and over all *Pongo* samples.

Step 2: Of the produced set of candidate SNPs, we identified high-confidence SNPs using the VQSR procedure implemented in the GATK. The principle of the method is to develop an estimate of the relationship between various SNP call annotations (*e.g.*, total depth, mapping quality, strand bias, etc.) and the probability that a SNP is a true genetic variant. The model is determined adaptively based on a set of ‘true SNPs’ (*i.e.*, known variants) provided as input. Our ‘true SNPs’ set contained 5,600 high-confidence SNPs, which were independently identified by three different variant callers in a previous reduced-representation sequencing project [71]. We ran the *Variant Recalibrator* of the GATK separately for each of the three raw SNP VCFs to produce recalibration files based on the ‘true SNPs’ and a VQSR training set of SNPs. The VQSR training sets were derived separately for each of the three raw SNP VCF files and contained the top 20% SNPs with highest variant quality score after having applied hard quality filtering as described for the VCF files in the BQSR procedure.

We used the produced VQSR recalibration files to filter the three candidate SNP VCFs with the Apply Recalibration walker of the GATK setting the ‘--truth_sensitivity_filter_level’ to 99.8%. Finally, we combined all SNPs of the three VCF files passing this filter using the *Combine Variants* tool of the GATK, hence generating a master list of high-confidence SNP sites in the genus *Pongo*.

Step 3: We called the genotype of each study individual at the identified high-confidence SNP sites. We performed genotyping on the recalibrated BAM files in multi-sample mode for Bornean and Sumatran orangutans separately, producing one SNP VCF file per island.

Finally, we only retained positions with high genome mappability, *i.e.*, genomic positions within a uniquely mappable 100-mers (up to 4 mismatches allowed), as identified with the GEM-mappability module from the GEM library build [76]. This mappability mask excludes genomic regions in the orangutan reference genome that are duplicated and therefore tend to produce ambiguous mappings, which can lead to unreliable genotype calling. Furthermore, we aimed to reduce spurious male heterozygous genotype calls on the X chromosome due to *UnifiedGenotyper* assuming diploidy of the entire genome. We determined the male-to-female ratios (M/F) of mean observed heterozygosity (H_o) and sequence coverage in non-overlapping 20-kb windows along the X chromosome across both islands. We obtained a list of X-chromosomal windows where M/F of H_o was above the 85%-quantile or M/F coverage was above the 95%-quantile, resulting in 1255 20-kb windows requiring exclusion. We then repeated step 3 of the genotype calling pipeline on the X chromosome for the male samples setting the argument ‘-ploidy’ of *UnifiedGenotyper* to 1 to specify the correct hemizygous state of the X chromosome in males. We subsequently masked all X-chromosomal positions within the spurious 20-kb windows in both male and female samples.

In total, we discovered 30,640,634 SNPs among all 37 individuals, which represent the most comprehensive catalogue of genetic diversity across the genus *Pongo* to date.

QUANTIFICATION AND STATISTICAL ANALYSIS

Recombination map estimation

We generated recombination maps for Bornean and Sumatran orangutans using the LDhat v2.2a software [77], following Auton et al. [78]. We used a high-quality subset of genotype data from the original SNP-calling dataset for the recombination map estimation for each island separately. Only biallelic, non-missing and polymorphic SNPs were used. Filtered genotype data were split into windows of 5,000 SNPs with an overlap of 100 SNPs at each side.

We ran the program *Interval* of the LDhat package for 60 million iterations, using a block penalty of 5, with the first 20 million iterations discarded as a burn-in. A sample was taken from the MCMC chain every 40,000 iterations, and a point estimate of the recombination rate between each SNP was obtained as the mean across samples. We joined the rate estimates for each window at the midpoint of the overlapping regions and estimated *theta per site* for each window using the finite-site version of the Watterson's estimate, as described in Auton & McVean [77].

We tested the robustness of the method with regards to the observed genome-wide variation of *theta* by contrasting recombination rate estimates using window-specific and chromosomal-average *thetas*. *Thetas* twice as large that the genome average produced very similar $4N_e r$ (*rho*) estimates. Because of this, a single genome-wide average of *theta per site* was used for all the windows (Sumatra: $\theta_w = 0.001917$, Borneo: $\theta_w = 0.001309$). We then applied additional filters following Auton et al. [78]. SNP intervals larger than 50 kb, or *rho* estimates larger than 100, were set to zero and the 100 surrounding SNP intervals (± 50 intervals) were set to zero recombination rate. A total of 1,000 SNP intervals were found to have $\rho > 100$ for *P. abelii*, and 703 for *P. pygmaeus*. In addition, 32 gaps (> 50 kb) were identified for *P. abelii*, and 47 gaps for *P. pygmaeus*. After applying the ± 50 interval criteria, a total of 7,424 SNP intervals were zeroed for *P. abelii*, and 15,694 for *P. pygmaeus*.

Haplotype phasing

We phased the genotype data from Bornean and Sumatran orangutans using a read aware statistical phasing approach implemented in SHAPEIT v2.0 [79, 80]. This allowed us to obtain good phasing accuracy despite our relatively low sample sizes by using phasing information contained in the paired-end sequencing reads to support the statistical phasing procedure. We used a high-quality subset of genotype data from the original SNP-calling dataset containing only biallelic and polymorphic SNPs. We first ran the program extractPIRs to extract phase informative reads (PIR) from the filtered BAM files. In a second step, we ran SHAPEIT in read aware phasing mode using the following parameters: 200 conditional states, 10 burnin iterations, 10 pruning iterations, 50 main iterations, and a window size of 0.5 Mb. Additionally, we provided two species-specific recombination maps (estimated with LDhat) and the PIR files obtained in the first step to the program.

SHAPEIT uses a recombination map expressed in cM/Mb, therefore it was necessary to convert the LDhat-based ρ estimates to cM/Mb units ($\rho=4N_e r$). Accordingly, we estimated island-specific effective population sizes using the Watterson's estimator of θ (Sumatra: $N_e[\theta_w]=41,000$, Borneo: $N_e[\theta_w]=27,000$) and applied these to the recombination map conversion. The most likely pair of haplotypes for each individual were retrieved from the haplotype graphs, and recoded into VCF file format.

Individual heterozygosity and inbreeding

We determined the extent of inbreeding for each individual by a genome-wide heterozygosity scan in sliding windows of 1 Mb, using a step size of 200 kb. We detected an excess of windows with very low heterozygosity in the density plots, pointing to some extent of recent inbreeding. To estimate the cutoff values of heterozygosity for the calculation of inbreeding coefficients, we calculated heterozygosity thresholds for each island according to the 5th-percentile of the genome-wide distribution of heterozygosities (Borneo: 1.0×10^{-4} heterozygote sites per bp; Sumatra: 1.3×10^{-4}). Neighboring regions with heterozygosities below the cutoff value were merged to determine the extent of runs of homozygosity (ROH). Based on the number and size of ROHs, we estimated the percentage of the genome that is autozygous, which is a good measure of inbreeding [81]. We choose 1 Mb as window size for the calculation of heterozygosities based on previous studies identifying regions smaller than 0.5 Mb as the result of background relatedness, and tracts larger than 1.6 Mb as evidence of recent parental relatedness [82].

Sex-specific genomic data: mitogenomes and Y chromosomes

We produced complete mitochondrial genome (mitogenome) sequences for all study individuals. We first created a consensus reference sequence from 13 Sanger-sequenced mitogenomes representing almost all major genetic clusters of extant orangutans using BioEdit v7.2.0. [83]. The Sanger-sequenced mitogenomes were generated via 19 PCRs with product sizes of 1.0–1.2 kb and an overlap of 100–300 bp (Table S6) following described methods [84]. PCR conditions for all amplifications were identical and comprised a pre-denaturation step at 94°C for 2 minutes, followed by 40 cycles each with denaturation at 94°C for 1 minute, annealing at 52°C for 1 minute, and extension at 72°C for 1.5 minutes. At the end, we added a final extension step at 72°C for 5 minutes. PCR products were checked on 1% agarose gels, excised from the gel and after purification with the Qiagen Gel Extraction Kit, sequenced on an ABI 3130xL sequencer using the BigDye Terminator Cycle Sequencing kit (Applied Biosystems) in both directions using the amplification primers.

We individually mapped Illumina whole-genome sequencing reads of all 37 study individuals (Table S4) to the consensus mitochondrial reference sequence using NovoAlign v3.02. (NovoCraft), which can accurately handle reference sequences with ambiguous bases. This procedure prevented biased

short read mapping due to common population-specific mutations. For each individual, we generated a FASTA sequence for the mitogenome with the *mpileup* pipeline of SAMtools. We only considered bases with both mapping and base Phred quality scores ≥ 30 and required all positions to be covered between 100 and 2000 times. Finally, we visually checked the sequence alignment of all individuals in BioEdit and manually removed indels and poorly aligned positions and excluded the D-loop to account for sequencing and alignment errors in those regions which might inflate estimates of mtDNA diversity. In total, we identified 1,512 SNPs among all 50 individuals.

We thoroughly investigated the literature for the potential occurrence of nuclear insertions of mtDNA (numts) in the genus *Pongo*, given that this has been a concern in closely related gorillas (*Gorilla* spp.) [85]. There was no indication of numts in the genus *Pongo*, which is in line with our own previous observations [28, 52, 53]. Numts also seem unlikely given our high minimal sequence depth threshold.

We developed a comprehensive bioinformatics strategy to extract sequences from the male-specific region of the Y chromosome (MSY) from whole-genome sequencing data. We expect the principle of our bioinformatics strategy to be applicable to mammalian species in general if the taxon under investigation is in phylogenetic proximity to one for which a Y-chromosomal reference sequence is present or will be made available. Like for most mammals, there is currently no reference Y chromosome for orangutans. Therefore, we had to rely on a reference assembly of a related species (*i.e.*, humans) for sequence read mapping. Despite the ~ 18 million years divergence between humans (*Homo* spp.) and orangutans [51, 86], we obtained a high number of MSY sequences. The impact of varying Y chromosome structure among species [87, 88] on sequence read mappability might have been reduced because we exclusively targeted X-degenerate regions. Hughes et al. [89] showed for human and chimpanzees that although less than 50% of ampliconic sequences have a homologous counterpart in the other species, over 90% of the X-degenerate sequences hold such a counterpart.

We applied several filters to ensure male-specificity and single-copy status of the generated MSY sequences. (i) We simultaneously mapped sequencing reads to the whole orangutan reference genome *PonAbe2* [50] and not just the human reference Y chromosome, reducing spurious mapping of autosomal reads to the Y chromosome and allowing subsequent identification of reads that also aligned to the X or autosomal chromosomes. (ii) We exclusively accepted reads that mapped in a proper pair, *i.e.*, where both read mates mapped to the Y chromosome, which considerably increased confidence in Y-specific mapping. (iii) We also mapped whole-genome sequencing reads of 23 orangutan females to the human Y reference chromosome and excluded all reference positions where female reads had mapped from the male Y sequence data. (iv) To exclude potential repetitive regions, we filtered non-uniquely mapped reads as well as positions with sequence coverage greater than two times the median coverage for each individual, as extensive coverage can be indicative for repetitive regions which might appear as collapsed regions on the Y reference chromosome. (v) To ensure that we only targeted unique,

single-copy MSY regions, we exclusively retained reads mapping to four well-established X-degenerate regions of the MSY in humans [90].

Our bioinformatics strategy consisted of the following detailed steps. First, we created a new reference sequence (*PonAbe2_humanY*) by manually adding the human reference Y chromosome (*GRCh37*) to the orangutan reference genome *PonAbe2* [50]. We then used BWA-MEM v0.7.5. [73] to map Illumina whole-genome short reads from 36 orangutans (13 males and 23 females) to this new reference sequence. We mapped reads for each individual separately in paired-end mode and with default settings. To reduce output file size, we removed unmapped reads on the fly using SAMtools v0.1.19 [91]. Picard v1.101 was used to add read groups and sort the BAM files. We then extracted all reads which mapped to the Y chromosome using SAMtools and marked read duplicates with Picard.

We used the GATK [74, 75] to perform local realignment around indels and filtered out duplicated reads, bad read mates, reads with mapping quality zero and reads which mapped ambiguously. We called genotypes at all sequenced sites with the *Unified Genotyper* of the GATK, applying the output mode 'EMIT_ALL_CONFIDENT_SITES'. We called genotypes in multi-sample mode (females and males separately, sample-ploidy was set to 1), producing one genomic VCF file for each sex. We only accepted bases/reads for genotype calling if they had Phred quality scores ≥ 30 .

From the VCF file of the females, we generated a 'nonspec' list with the coordinates of all sites with coverage in more than one female (minimal sequence depth 2x), as these sites most likely were located in pseudoautosomal or ampliconic regions, *i.e.*, share similarity with the X or autosomal chromosomes. To ensure Y-specificity, we removed all sites of the 'nonspec' list from the VCF file of the males with VCFtools v0.1.12b. [92].

Finally, we used GATK to extract sequences of four well-established X-degenerate regions of the MSY in humans (14,170,438–15,795,786; 16,470,614–17,686,473; 18,837,846–19,267,356; 21,332,221–21,916,158 on the human reference Y chromosome assembly GRCh37/hg19)[90]. To be conservative, we chose regions which were longer than 1 Mb in humans and disregarded the first and last 300 kb of each region to account for potential uncertainties regarding region boundaries, leaving us with 3,854,654 bp in total. We exclusively retained genotype calls that were covered by a minimum of two reads and had a maximum of twice the individual mean coverage, resulting in 2,825,271 bp of MSY sequences among the 13 orangutan males. As expected, individual mean MSY sequence depth was about half (average: 54.4%) of that recorded for the autosomes, and ranged from 2.79–16.62x. For analyses, we only kept sites without missing data, *i.e.*, with a genotype in all study males. Because genomes of some individuals had been sequenced to only low coverage (~5–7x) [50], this left us with 673,165 bp of MSY sequences. We identified 1,317 SNPs among the 13 males, corresponding to a SNP density of 1 SNP every 511 bp.

We constructed phylogenetic trees and estimated divergence dates for mitogenome and MSY sequences using the Bayesian Markov chain Monte Carlo (MCMC) method implemented in BEAST v1.8.0. [58]. To determine the most suitable nucleotide substitution model, we conducted model selection with jModelTest v2.1.4. [60]. Based on the Akaike information criterion (AIC) and corrected AIC, we selected the GTR+I substitution model [93] for mitogenomes and the TVM+I+G model [94] for MSY sequences.

The mitogenome tree was rooted with a human and a central chimpanzee sequence from GenBank (accession numbers: GQ983109.1 and HN068590.1), the MSY tree with the human reference sequence *hg19*. We estimated divergence dates under a relaxed molecular clock model with uncorrelated lognormally distributed branch-specific substitution rates [95]. The prior distribution of node ages was generated under a birth-death speciation process [96]. We used fossil based divergence estimates to calibrate the molecular clock by defining a normal prior distribution for certain node ages. For mitogenomes, we applied two calibration points, *i.e.*, the *Pan-Homo* divergence with a mean age of 6.5 Ma and a standard deviation of 0.3 Ma [97, 98] and the Ponginae-Homininae divergence with a mean age of 18.3 Ma and a larger standard deviation of 3.0 Ma [86], which accounts for the uncertainty in the divergence date [99]. For MSY sequences, we used the Ponginae-Homininae divergence for calibration. We performed four independent BEAST runs for 30 million generations each for mitogenomes, with parameter sampling every 1,000 generations, and for 200 million generations each with parameter sampling every 2,000 generations for MSY sequences. We used Tracer v1.6 [100] to examine run convergence, aiming for an effective sample size of at least 1000 for all parameters. We discarded the first 20% of samples as burn-in and combined the remaining samples of each run with LogCombiner v1.8.0. [58]. Maximum clade credibility trees were drawn with TreeAnnotator v1.8.0. [58] and trees visualized in FigTree v1.4.0. [101] and MEGA v6.06. [102].

Autosomal genetic diversity and population structure

For all subsequent population genetic analyses, we assumed an autosomal mutation rate (μ) of 1.5×10^{-8} per base pair per generation, based on estimates obtained for the present-day mutation rates in humans and chimpanzees, derived primarily from de novo sequencing comparisons of parent-offspring trios but also other evidence [103-106]. There is good reason to believe that the mutation rate in orangutans is similar to that in other great apes, given the very similar branch lengths from outgroups such as gibbon and macaque to each species [107]. We assumed a generation time of 25 years [108].

We identified patterns of population structure in the autosomal genome by principal component analysis (PCA) of biallelic SNPs using the function ‘prcomp’ in R v3.2.2 [109]. Three separate analyses were performed: one within each island and one including all study individuals. For each sample set, we excluded all genotypes from the SNP VCF files that were covered by less than five reads and only retained SNPs with a genotype call in all individuals after this filter. Furthermore, we removed SNPs

with more than two alleles and monomorphic SNPs in the particular sample set. This restrictive filtering left us with 3,006,895 SNPs for the analysis of all study individuals, 5,838,796 SNPs for PCA within Bornean orangutans and 4,808,077 SNPs for PCA within Sumatran orangutans.

We inferred individual ancestries of orangutans using ADMIXTURE v1.23 [110]. We randomly sampled one million sites from the original VCF files and filtered this subset by excluding sites with missing genotypes or with a minor allele frequency less than 0.05. We further reduced the number of sites to 272,907 by applying a linkage disequilibrium (LD) pruning filter using PLINK v1.90b3q (–indep-pairwise 50 5 0.5) [111]. ADMIXTURE was run 20 times at all K values between 1 and 10. Among those runs with a difference to the lowest observed cross validation (CV) error of less than 0.1 units, we reported the replicate with the highest biological meaning, *i.e.*, runs that resolved substructure among different sampling areas rather than identifying clusters within sampling areas.

For subsequent analyses, we defined seven distinct populations based on the results of the PCA and ADMIXTURE analyses: three on Sumatra (Northeast Alas comprising North Aceh and Langkat regions, West Alas, and Batang Toru) and four on Borneo (East Kalimantan, Sarawak, Kinabatangan comprising North and South Kinabatangan, and Central/West Kalimantan comprising Central and West Kalimantan). Even though individuals from North and South Kinabatangan could be clearly distinguished in the PCA and ADMIXTURE analysis, we decided to pool the two Kinabatangan populations due to their low samples sizes ($n = 2$). This can be justified as data from the mitochondrial genome showed that they started to diverge only recently (~ 40 ka).

Ancestral gene flow between orangutan populations

We used D-statistics to assess gene flow between orangutan species, testing all three possible phylogenetic relationships among *P. abelii*, *P. tapanuliensis*, and *P. pygmaeus*. We extracted genotype data from the two individuals per population with the highest sequencing coverage and included two human genome sequences as outgroup (SRA sample accession: ERS007255 and ERS007266). We calculated D-statistics for all combinations of populations involving the three species using the qpDstat program of the ADMIXTOOLS package v4.1 and assessed significance using the block jackknife procedure implemented in ADMIXTOOLS.

To explore temporal patterns of gene flow between orangutan populations, we applied the multiple sequential Markovian coalescent (MSMC2) model [112]. The rate of coalescence of between-population haplotype pairs was compared to the within-population coalescence rate of haplotype pairs from the same population to obtain the relative cross-coalescence rate (RCCR) through time. A RCCR close to 1 indicates extensive gene flow between populations, while a ratio close to 0 indicates complete genetic isolation.

We used the phased whole-genome data for the relative cross-coalescence rate analysis. To avoid coverage-related issues, we selected the individual with the highest sequencing coverage for each population. We further excluded sites with an individual sequencing coverage less than 5x, a mean mapping quality less than 20, or sites with low mappability based on the mappability mask.

We ran MSMC2 for all pairs of populations, using a single individual (*i.e.*, two haplotypes) per population. For each population pair, we performed three individual MSMC2 runs, using the default time discretization parameters: within population 1 (two haplotypes; -I 0,1), within population 2 (two haplotypes; -I 2,3), and between populations (four haplotypes; -I 0,1,2,3 -P 0,0,1,1). We then used the combineCrossCoal.py Python script of the MSMC2 package to combine the outputs of the three runs into a combined output file.

As the sequencing coverage of the best Batang Toru individual was substantially lower compared to individuals from other populations (~17x vs. ~23–27x, Table S4), we also assessed whether different sequencing coverage was negatively affecting the relative cross-coalescence rate results. To achieve this, we repeated the analysis using individuals with similar coverage as the Batang Toru individual (~16–21x). The results were highly consistent with the output from the runs with the highest-coverage individuals, indicating that the relative cross-coalescent rate analysis was robust to differences in sequencing coverage in our data set.

Approximate Bayesian Computation (ABC)

To gain insights into the colonization history of the Sundaland region by orangutans and obtain parameter estimates of key aspects of their demographic history, we applied a model-based ABC framework [31]. For this, we sampled a total of 3,000 independent sequence loci of 2 kb each, following the recommendations in Robinson et al. [113]. Loci were sampled randomly from non-coding regions of the genome, with a minimum distance of 50 kb between loci to minimize the effects of linkage. Since the coalescent simulations underlying ABC inference assume neutrality, we excluded loci located within 10 kb of any exonic region defined in the *Pongo abelii* Ensembl gene annotation release 78, as well as loci on the X chromosome and the mitochondrial genome, which would exhibit reduced N_e as compared to the autosomal regions.

For all ABC-based modelling, we defined three metapopulations for the calculation of summary statistics: Sumatran populations north of Lake Toba (NT), the Sumatran population of Batang Toru south of Lake Toba (ST), as well as all Bornean populations (BO). For each metapopulation as well as over all metapopulations combined, we calculated the first four moments over all loci for the following summary statistics: nucleotide diversity (π), Watterson's theta, and Tajima's D. Furthermore, for each of the three pairwise comparisons between metapopulations, we calculated the first four moments over loci of the number of segregating sites, proportions of shared and fixed polymorphism, average sequence divergence (d_{XY}), and Φ_{ST} [114]. To avoid potential problems with unreliable phasing, we

only used summary statistics that do not require phased sequence data. This resulted in a total of 108 summary statistics used in the ABC analyses. For each locus, we extracted genotype data of a total of 22 individuals (5 Northeast Alas, 5 West Alas, 2 Batang Toru, 4 Central/West Kalimantan, 2 East Kalimantan, 2 Sarawak, 2 Kinabatangan) by selecting the individuals with the highest sequence coverage for a given locus. Additionally, we recorded the positions of missing data for each locus and individual and coded genotypes as ‘missing’ in the simulated data if mutations fell within the range of missing data in the observed data.

In a first step, we used a model testing framework to infer the most likely sequence of population splits in the colonization history of orangutans. For this, we designed four models representing potential colonization patterns into Sundaland (Figure 3A). We assumed a simplified population structure with three distinct, random mating units composed of NT, ST, and BO metapopulations as described above. We simulated 4×10^6 data sets for each model using the coalescent simulator ms [115]. Since we obtained a large number of summary statistics, we used a partial least squares discriminant analysis (PLS-DA) to extract the orthogonal components of the summary statistics that are most informative to discriminate between the four competing models using the ‘plsda’ function of the R package ‘mixOmics’ v5.2.0 [116] in R version 3.2.2 [109]. For model testing, we used the R package ‘abc’ v2.1 [117] to perform a multinomial logistic regression on the PLS transformed simulated and observed summary statistics, using a tolerance level of 0.05% (8,000 simulations closest to the observed data). To find the optimal number of PLS components for model selection, we performed cross-validations with 200 randomly chosen sets of summary statistics for each model and assessed model misspecification rates when using 10, 12, 15, 18, and 20 components.

We found that using the first 18 PLS components resulted in the lowest model misspecification rate. However, our model testing approach lacked power to reliably differentiate between pairs of models with the same underlying species tree (*i.e.*, model 1a vs. model 1b and model 2a vs. model 2b in Figure 3A), as evidenced by a high model misspecification rate of 47.63% across all four models. In order to increase discrimination power with a new set of optimized PLS components, we therefore repeated the PLS-DA and multinomial logistic regression with the two best-fitting models (model 1a vs. model 1b). This resulted in a substantially lower model misspecification rate (36.00%). Moreover, no model misassignment occurred with a posterior probability equal or higher than the observed value (0.976), indicating a high confidence in the selected model (model 1a).

After establishing the order of population split events, we were interested in parameter estimates of different aspects of the orangutan demographic history. For this, we applied a more complex model that included additional population structure in NT and BO, as well as recent population size changes (Figure 3B). The design of this model was informed by (i) PCA and ADMIXTURE analyses (Figs. 2B and 2C), (ii) MSMC2 analyses (Figure 3C), and (iii) previous demographic modeling using more limited sets of genetic markers [57]. For parameter estimation, we performed a total of 1×10^8 simulations

as described above. Model parameterization and parameter prior distributions are shown in Table S5. We used 100,000 random simulations to extract the orthogonal components of the summary statistics that maximize the covariance matrix between summary statistics and model parameters using the ‘plsr’ function of the R package ‘pls’ v2.5-0 [118]. We defined the optimal number of partial least squares (PLS) components based on the drop in the root mean squared error for each parameter with the inclusion of additional PLS components [119]. After transforming both the simulated and observed summary statistics with the loadings of the extracted PLS components, we performed ABC-GLM post-sampling regression [120] on the simulations with the smallest Euclidean distance to the observed summary statistics using ABCtoolbox v2.0 [121]. To find the optimal proportion of retained simulations, we assessed the root-mean-integrated-squared error of the parameter posterior distributions based on 1,000 pseudo-observed data sets (pods) randomly chosen from the simulated data. We found that varying the tolerance level had little impact on the accuracy of the posterior distributions and therefore used a tolerance level of 0.00002 (equaling 2,000 simulations) for parameter estimation.

To assess the goodness of fit of our demographic model, we calculated the marginal density and the probability of the observed data under the general linear model (GLM) used for the post-sampling regression with ABCtoolbox [120]. A low probability of the observed data under the GLM indicates that the observed data is unlikely to have been generated under the inferred GLM, implying a bad model fit. We obtained a p-value of 0.14, showing that our complex demographic model is well able to reproduce the observed data. Additionally, we visualized the coverage of summary statistics generated under the demographic model relative to the observed data by plotting the first 12 principal components of the simulated and observed data. For this, we randomly selected 100,000 simulations and extracted PCA components using the ‘prcomp’ function in R. The observed data fell well within the range of simulated summary statistics for all 12 components. Furthermore, we checked for biased posterior distributions by producing 1,000 pods with parameter values drawn from the prior distributions. For each pods, we determined the quantile of the estimated posterior distribution within which the true parameter values fell and used a Kolmogorov-Smirnov in R to test the resulting distribution of posterior quantiles for uniformity. Deviations from uniformity indicate biased posterior distributions [122] and the corresponding parameter estimates should be treated with caution. As expected from complex demographic models, multiple parameters showed significant deviations from uniformity after sequential Bonferroni correction [123]. However, in most of these distributions, data points were overrepresented in the center of the histogram, which indicates that posterior distributions were estimated too conservatively.

G-PhoCS analysis

We used the full-likelihood approach implemented in G-PhoCS v1.2.3 [124] to compare different models of population splitting with gene flow and to estimate parameters of the best-fitting model. Due to computational constraints, we limited our data set to eight individuals with good geographic coverage of the extant orangutan distribution (1 Northeast Alas, 1 West Alas, 2 Batang Toru, 2 Central/West Kalimantan, 1 East Kalimantan, 1 Kinabatangan). We sampled 1-kb loci across the autosomal genome, ensuring a minimum distance of 50 kb among loci to minimize linkage. To reduce the impact of natural selection, we excluded loci located within 1 kb of any exonic region defined in the *Pongo abelii* Ensembl gene annotation release 78. We coded sites as missing based on the following filter criteria: low mappability, mean mapping quality less than 20, and individual coverage less than 5x. Sites without at least one valid genotype per species were excluded completely. We only retained loci with at least 700 bp of sites with data, resulting in a total of 23,380 loci for which we extracted genotype information for the eight selected individuals.

We compared models with the three different possible underlying population trees in a three taxon setting (Borneo, Sumatra north of Lake Toba, and Batang Toru). We performed 16 independent G-PhoCS runs for each model, running the MCMC algorithm for 300,000 iterations, discarding the first 100,000 iterations as burn-in and sampling every 11th iteration thereafter. The first 10,000 iterations were used to automatically adjust the MCMC finetune parameters, aiming for an acceptance rate of the MCMC algorithm of 30–40%. We merged the resulting output files of independent runs and analysed them with Tracer v1.6 [100] to ensure convergence among runs. We then used the model comparison based on the Akaike information criterion through MCMC (AICM) [125, 126] implemented in Tracer to assess the relative fit of the three competing models.

In agreement with the ABC analyses, the model positing the deepest split between Sumatra north of Lake Toba and Batang Toru, followed by a split between south of Lake Toba and Borneo, showed a much better fit to the data compared to the two other splitting patterns. Independent replicates of the same model produced highly consistent posterior distributions, indicating convergence of the MCMC algorithm. All parameters of the best-fitting model were estimated with high precision, as shown by the small 95%-highest posterior density ranges (Table S5). Compared to the estimates from the ABC analysis, G-PhoCS resulted in more recent divergence time estimates for both the NT/(BO,ST) and BO/ST splits. This discrepancy might be caused by hypermutable CpG sites, which likely violate certain assumptions of the G-PhoCS model [124]. We could not exclude CpG sites in our analysis due to the absence of a suitable outgroup for calibration. Instead, we had to rely on a fixed genome-wide mutation rate, which includes hypervariable CpG sites. An alternative explanation could be a likely bias in the G-PhoCS results due to the restriction to a highly simplified demographic model as compared to our ABC analyses; G-PhoCS assumes constant effective population sizes and migration rates in between

population splits. However, this assumption is most likely violated in orangutans, as shown by the results of our ABC analysis (Figure 3B, Table S5).

Cranial, dental, and mandibular morphology

We evaluated five qualitative and 44 quantitative cranial, dental, and mandibular variables (Tables S1 and S2). We chose variables that had previously been used to describe and differentiate orangutan cranio-mandibular shape [61-63, 127-132]. Due to extensive dental wear of the Batang Toru specimen, we limited our comparisons with the Padang cave material to the breadth of the upper and lower canines, in addition to the length, breadth, and area (*i.e.*, breadth x length) of the lower first molar, all of which displayed a limited amount of wear. All measurements were taken by a single individual (AnN) in order to reduce observer bias.

We used both univariate and multivariate statistics to evaluate the Batang Toru specimen in relation to our comparative sample. As Batang Toru is only represented by a single sample, we first compared it to the interquartile range (IQR, defined as the range between the first and the third quartile) and the lower and upper inner fence ($\pm 1.5 \times \text{IQR}$) for each separate sample population, using traditional methods for evaluating outliers [133]. This allowed us to evaluate the Batang Toru specimen's distance and direction from the central tendency of our sample orangutan populations. We also conducted univariate exact permutation tests for each morphological variable by removing a single sample for either the *P. abelii*, *P. pygmaeus*, or *P. p. palaeosumatrensis* sample populations and then comparing the linear distance to the mean of the remaining samples. This was done for each sample until all samples had a calculated value. A linear distance between the *P. tapanuliensis* sample and the *P. abelii*, *P. pygmaeus*, and *P. p. palaeosumatrensis* mean values (*i.e.*, the test statistics) was then calculated and compared to the sample distributions detailed above. P-values represent the number of samples from the sample distribution that exceed the test statistic, divided by the total number of comparisons. In some cases, specimens did not preserve the measurements utilized in this study (*e.g.*, broken bone elements and/or missing/heavily worn teeth), and so were excluded from comparisons. Sample sizes for univariate comparisons of extant orangutan cranio-mandibular morphology are detailed in Table S1, whereas the sample sizes for the univariate comparisons of extant and fossil teeth are detailed in Table S2.

We also conducted a PCA on 26 of our 39 cranio-mandibular variables, on a subset of our extant orangutan sample, including *P. abelii* (n=8), *P. pygmaeus* (n=19), and the newly described *P. tapanuliensis* specimen. The choice of 26 variables allowed us to maximize sample size and avoid violating the assumptions of PCA [134]. A scree plot (using the *princomp* function from the base *stats* package in R [135]) indicated that seven principal components were sufficient to be extracted, based on the Kaiser criterion of eigenvalues at ≥ 1 [136]. Using the *principal* function from the *psych* R package [137], we ran a PCA on the correlation matrix of our 26 selected variables, extracting seven principal components with varimax rotation.

To highlight the multivariate uniqueness of *P. tapanuliensis*, we used the extracted PCs and calculated the Euclidean D^2 distance for each sample relative to the *P. abelii* and *P. pygmaeus* centroids. We grouped these distances into two distributions, referred to as the between species (*i.e.*, the distances of all *P. abelii* samples to the *P. pygmaeus* centroid plus all of the *P. pygmaeus* samples to the *P. abelii* centroid) and within species (*i.e.*, the distances of all *P. abelii* samples to the *P. abelii* centroid plus all of the *P. pygmaeus* samples to the *P. pygmaeus* centroid) distributions. We then compared the Euclidean D^2 distances of *P. tapanuliensis* to the *P. abelii* and *P. pygmaeus* centroids (*i.e.*, the test values), relative to the two aforementioned sample distributions. Exact permutation p-values for these results were calculated as the number of samples from the sample distribution that exceed the test statistic, divided by the total number of comparisons. All Euclidean D^2 distance were calculated in the base *stats* package in R [135].

Acoustic and behavioral analyses

We used both previously published [138-140] and newly collected data in our analyses of male long calls. The current study includes $n=130$ calls from $n=45$ adult males across 13 orangutan field sites. In addition to two individuals from Batang Toru, we sampled 14 individuals of *P. abelii* and 29 individuals of *P. pygmaeus*. Using our comparative sample, we evaluated 15 long call variables (Table S3). We chose variables and their definitions that had previously been described to differentiate orangutan male long calls [138, 139, 141].

We used both univariate and multivariate statistics to evaluate the Batang Toru specimen in relation to our comparative sample. As Batang Toru is only represented by two individuals, we compared the mean of these two sample points to the interquartile range (IQR) and the lower and upper inner fence ($\pm 1.5 \times \text{IQR}$) for each separate sample population [133]. As above, univariate exact permutation tests were conducted for each long call variable by removing a single sample for either the *P. abelii* or *P. pygmaeus* sample populations and then comparing the linear distance to the mean of the remaining samples. This was done for each sample until all samples had a calculated value. A linear distance between the average of the two *P. tapanuliensis* samples and the *P. abelii* or *P. pygmaeus* mean values (*i.e.*, the test statistics) was then calculated and compared to the sample distributions detailed above. P-values represent the number of samples from the sample distribution that exceed the test statistic, divided by the total number of comparisons. In some cases, not all acoustic variables were available for each individual. As such, sample sizes for univariate comparisons are detailed in Table S3.

Geological and ecological analyses

We evaluated five ecological variables, including the type and age of geological parent material, elevation, average temperature, and average rainfall, to highlight that the current ecological niche of *P.*

tapanuliensis is divergent relative to that of *P. abelii* and *P. pygmaeus*. For Sumatran populations, type and age of geological parent material were digitized from the land unit and soil map series of Sumatra [142-149]. No comparable geospatial data is available for Borneo, so we used previously published materials to more broadly characterize areas populated by orangutans [150]. To maintain consistency, elevation, average temperature, and average annual rainfall were collected from the WorldClim v. 1.4 bioclimatic variables dataset [151]. Using the digitized land unit/soil maps, we calculated the percentage of Sumatran orangutan distribution [152] classified into four classes for each type (*e.g.*, igneous, metamorphic, sedimentary, and other rock [*i.e.*, land units with a mixture of rock types]) and age (*e.g.*, Pre-Cenozoic, Tertiary, Quaternary, and other [*i.e.*, land units with a mixture of ages]) of geological parent material. For the elevation and climatic variables, we created 1km x 1km sample point grids for each currently identified orangutan population in Borneo and Sumatra [152, 153], and sampled the three aforementioned WorldClim datasets.

DATA AND SOFTWARE AVAILABILITY

Raw sequence read data have been deposited into the European Nucleotide Archive (ENA; <http://www.ebi.ac.uk/ena>) under study accession number PRJEB19688. Mitochondrial and Y-chromosomal sequences are available from the Mendeley Data repository under ID code doi:10.17632/hv2r94yz5n.1.

KEY RESOURCES TABLE

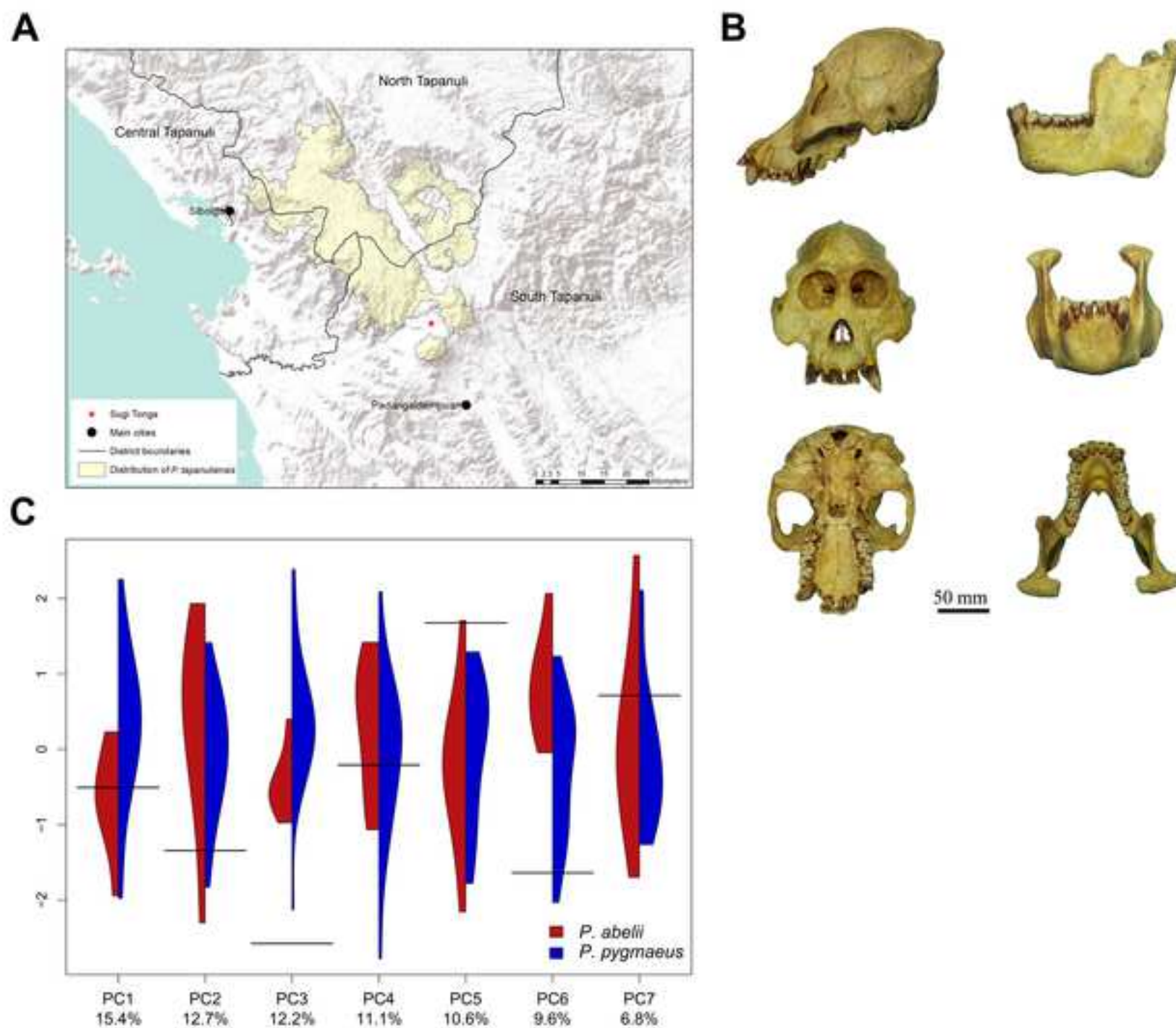
REAGENT or RESOURCE	SOURCE	IDENTIFIER
Biological Samples		
17 <i>Pongo</i> spp. whole blood samples	This paper	See Table S4
34 <i>Pongo</i> spp. cranial specimens	This paper	N/A
Chemicals, Peptides, and Recombinant Proteins		
Proteinase K (20 mg/ml)	Promega	Cat#V3021
Critical Commercial Assays		
Gentra Puregene Blood Kit	Qiagen	Cat#158467
Deposited Data		
<i>Pongo abelii</i> reference genome <i>ponAbe2</i>	[50]	http://genome.wustl.edu/genomes/detail/pongo-abelii/
<i>Pongo abelii</i> Ensembl gene annotation release 78	Ensembl	https://www.ensembl.org/Pongo_abelii/Info/Index
Human reference genome NCBI build 37, GRCh37	Genome Reference Consortium	http://www.ncbi.nlm.nih.gov/projects/genome/assembly/grc/human/
Whole-genome sequencing data of 5 <i>Pongo abelii</i>	[50]	SRA: PRJNA20869
Whole-genome sequencing data of 5 <i>Pongo pygmaeus</i>	[50]	SRA: PRJNA74653
Whole-genome sequencing data of 10 <i>Pongo</i> spp.	[51]	SRA: PRJNA189439
Whole-genome sequencing data of 17 <i>Pongo</i> spp.	This paper	ENA: PRJEB19688
Whole-genome sequencing data of 2 <i>Homo sapiens</i>	Human Genome Diversity Project	SRA: ERS007255 and ERS007266
13 <i>Pongo</i> MSY sequences	This paper	http://dx.doi.org/10.17632/hv2r94yz5n.1
50 <i>Pongo</i> mitochondrial genome sequences	This paper	http://dx.doi.org/10.17632/hv2r94yz5n.1
Pictures of paratypes	This paper	https://morphobank.org/index.php/Projects/ProjectOverview/project_id/2591
Additional supporting information and analyses	This paper	https://morphobank.org/index.php/Projects/ProjectOverview/project_id/2591
Oligonucleotides		
19 mitochondrial primer pairs	This paper	See Table S6
Software and Algorithms		
FastQC v0.10.1.	[72]	https://www.bioinformatics.babraham.ac.uk/projects/fastqc/
BWA v0.7.5	[73]	http://bio-bwa.sourceforge.net/
Picard Tools v1.101		http://broadinstitute.github.io/picard/

GATK v3.2.2.	[74, 75]	https://software.broadinstitute.org/gatk/
GEM library	[76]	http://algorithms.wtf/gem-library
LDhat v2.2a	[77]	https://github.com/auton1/LDhat
SHAPEIT v2.0	[79]	https://mathgen.stats.ox.ac.uk/genetics_software/shapeit/shapeit.html
BioEdit v7.2.0.	[154]	http://www.mbio.ncsu.edu/bioedit/page2.html
NovoAlign v3.02.	Novocraft	http://www.novocraft.com/products/novoalign/
SAMtools v0.1.19	[155]	http://www.htslib.org/
VCFtools v0.1.12b.	[156]	https://vcftools.github.io/index.html
BEAST v1.8.0.	[58]	http://beast.community/index.html
jModelTest v2.1.4.	[60]	https://github.com/ddarriba/jmodeltest2
Tracer v1.6		http://tree.bio.ed.ac.uk/software/tracer/
FigTree v1.4.0.		http://tree.bio.ed.ac.uk/software/figtree/
MEGA v6.06.	[102]	http://www.megasoftware.net/mega.php
R 3.2.2	[109]	https://www.r-project.org
ADMIXTURE v1.23	[110]	https://www.genetics.ucla.edu/software/admixture/index.html
PLINK v1.90b3q	[111]	https://www.cog-genomics.org/plink2
ADMIXTOOLS v4.1	[157]	https://github.com/DReichLab/AdmixTools
MSMC2	[112]	https://github.com/stschiff/msmc2
ms	[115]	http://home.uchicago.edu/rhudson1/source/mksamples.html
R package 'mixOmics' v5.2.0	[116]	https://www.rdocumentation.org/packages/mixOmics
R package 'abc' v2.1	[117]	https://cran.r-project.org/package=abc
R package 'pls' v2.5-0	[118]	https://cran.r-project.org/package=pls

ABCtoolbox v2.0	[121]	http://www.unifr.ch/biology/research/wegmann/wegmannsoft
G-PhoCS v1.2.3	[124]	http://compgen.cshl.edu/GPhoCS/
R package 'psych'	[137]	https://cran.r-project.org/package=psych
R package 'MASS'	[158]	https://cran.r-project.org/package=MASS

Figure 1

[Click here to download Figure Figure1.TIF](#)



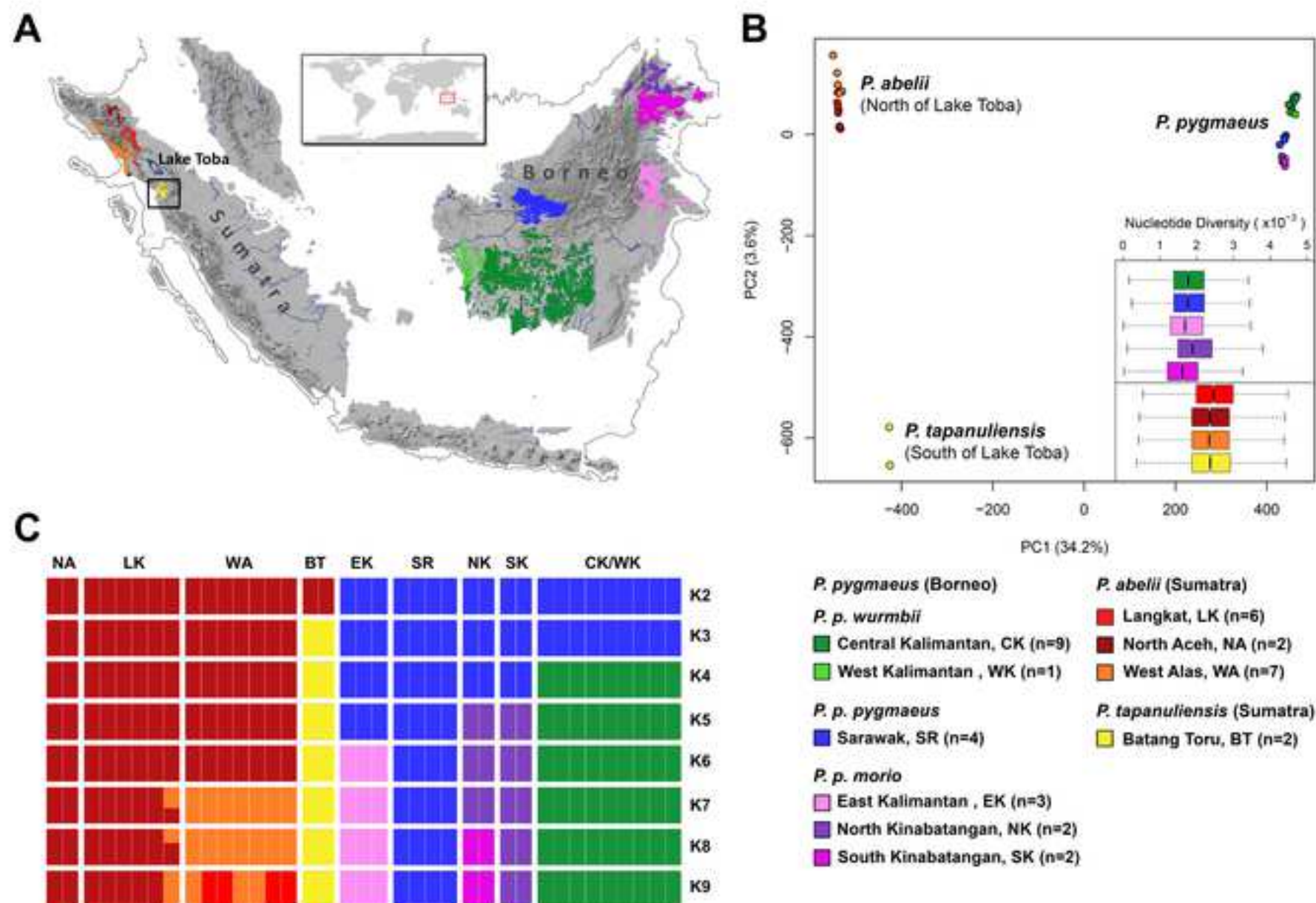
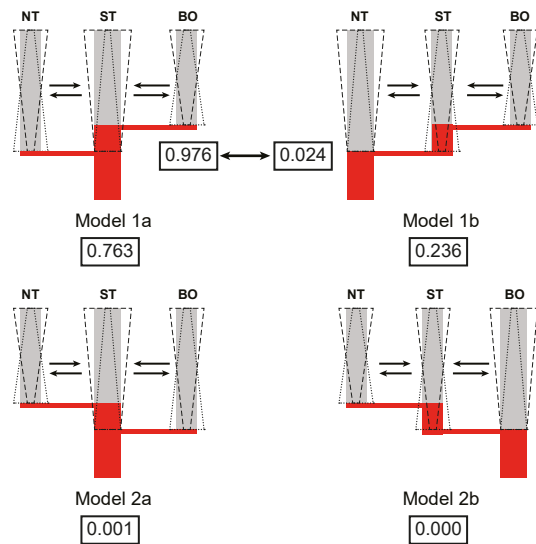
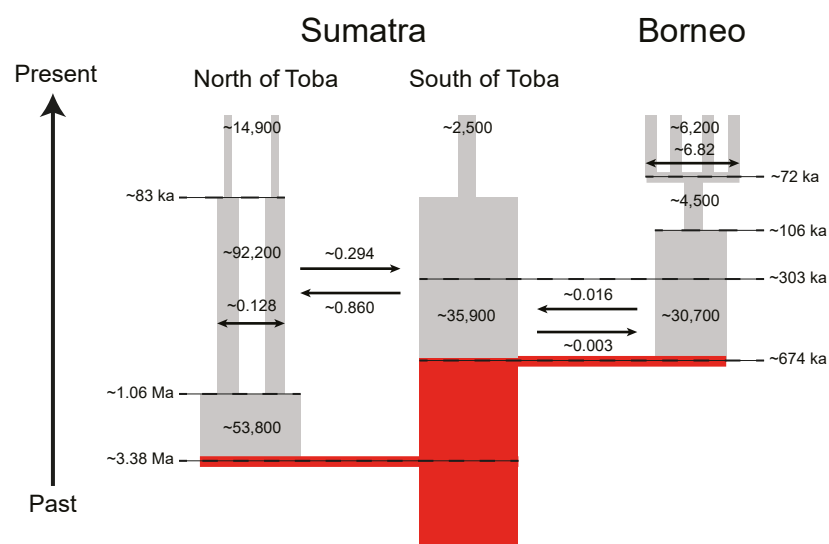


Figure 3

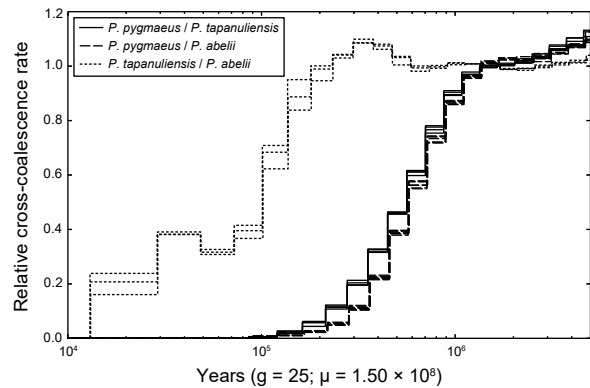
A



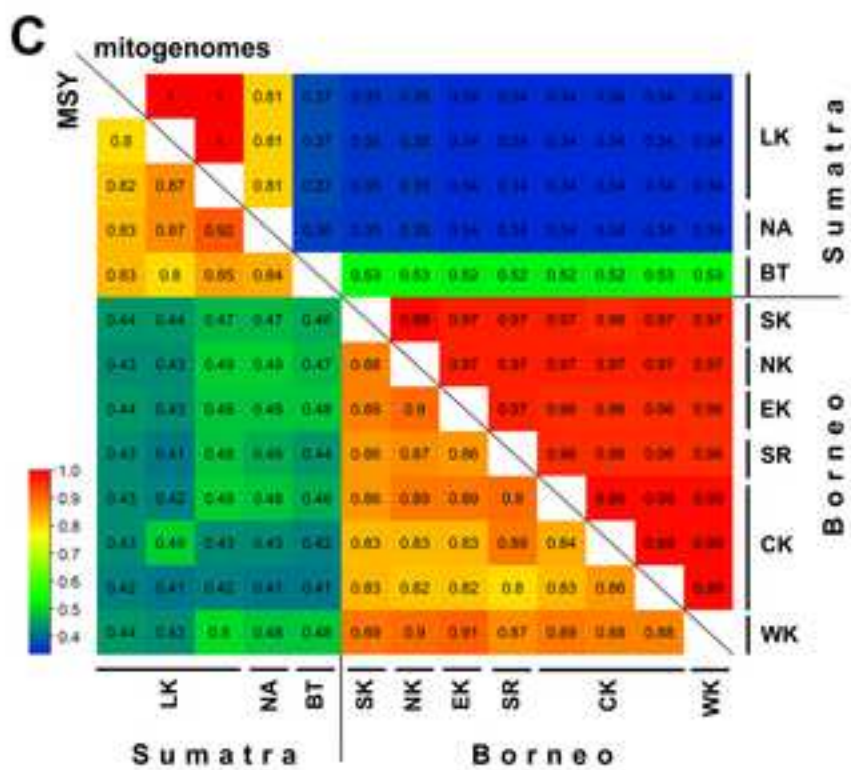
B



C



[Click here to download Figure Figure4.TIF](#)



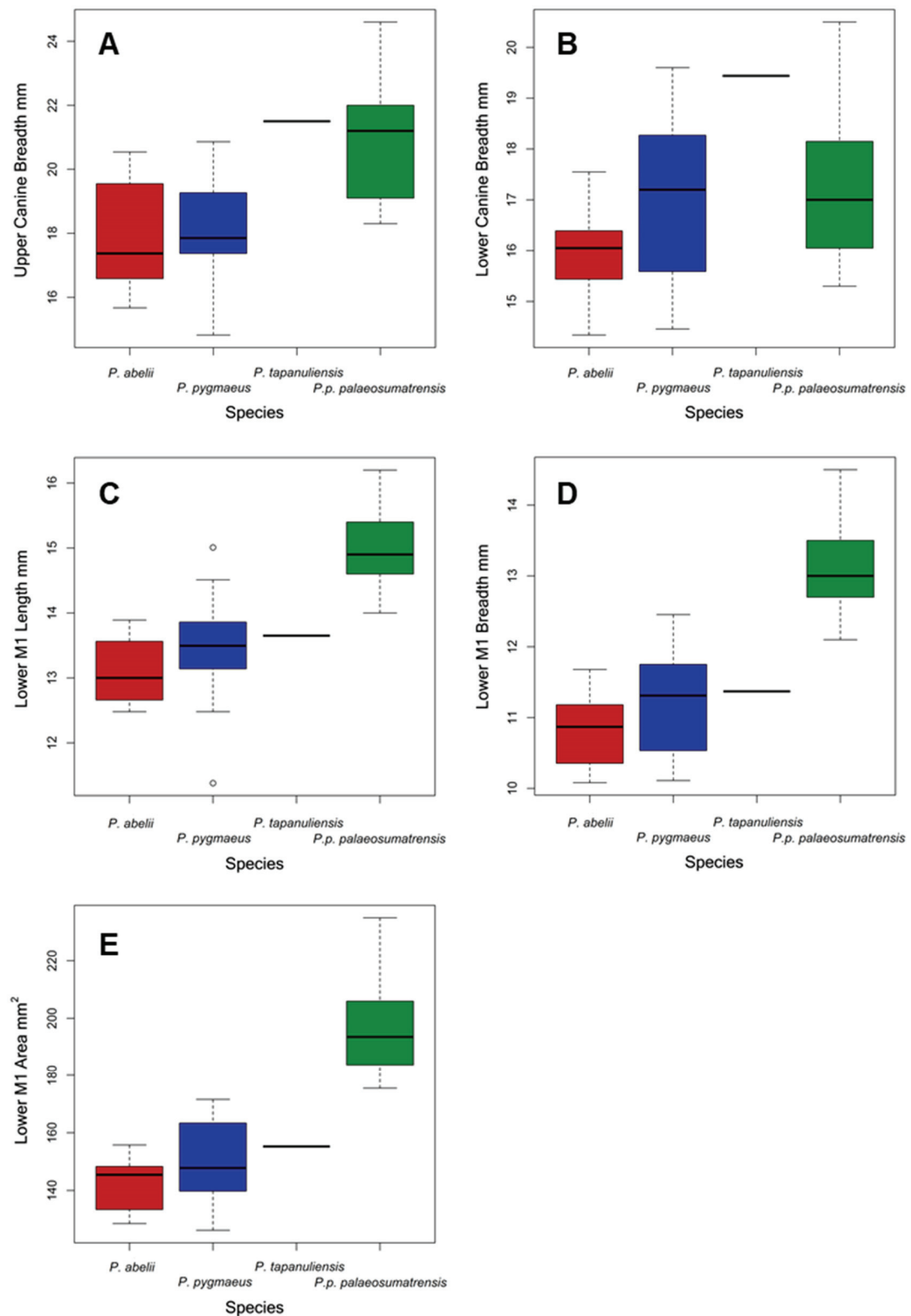


Figure S1. Comparisons of five dental variables across *P. abelii* (red), *P. pygmaeus* (blue), *P. tapanuliensis* (black horizontal line), and *P. p. palaeosumatrensis* (green). Related to Figure 1B. Variables include upper canine breadth (A), lower canine breadth (B), lower M1 length (C), lower M1 breadth (D), and lower M1 area (E). For each boxplot, the middle line is the median value of the

distribution, with the box representing the first (lower extreme) and third (upper extreme) quartile values (*i.e.*, the interquartile range [IQR]), and the whiskers representing the lower and upper extreme values that are within 1.5 x IQR of the first and third quartile values. Exact permutation analyses suggested that *P. tapanuliensis* could be differentiated statistically from the *P. abelii* mean for both the upper (p-value<0.001) and lower canine breadths (p-value<0.001) and from the *P. 'pygmaeus'* *palaeosumatrensis* mean for lower M₁ length (p-value<0.001), breadth (p-value<0.001), and area (p-value<0.001). *P. tapanuliensis* could not be differentiated statistically from the *P. pygmaeus* mean for any of the five dental measures.

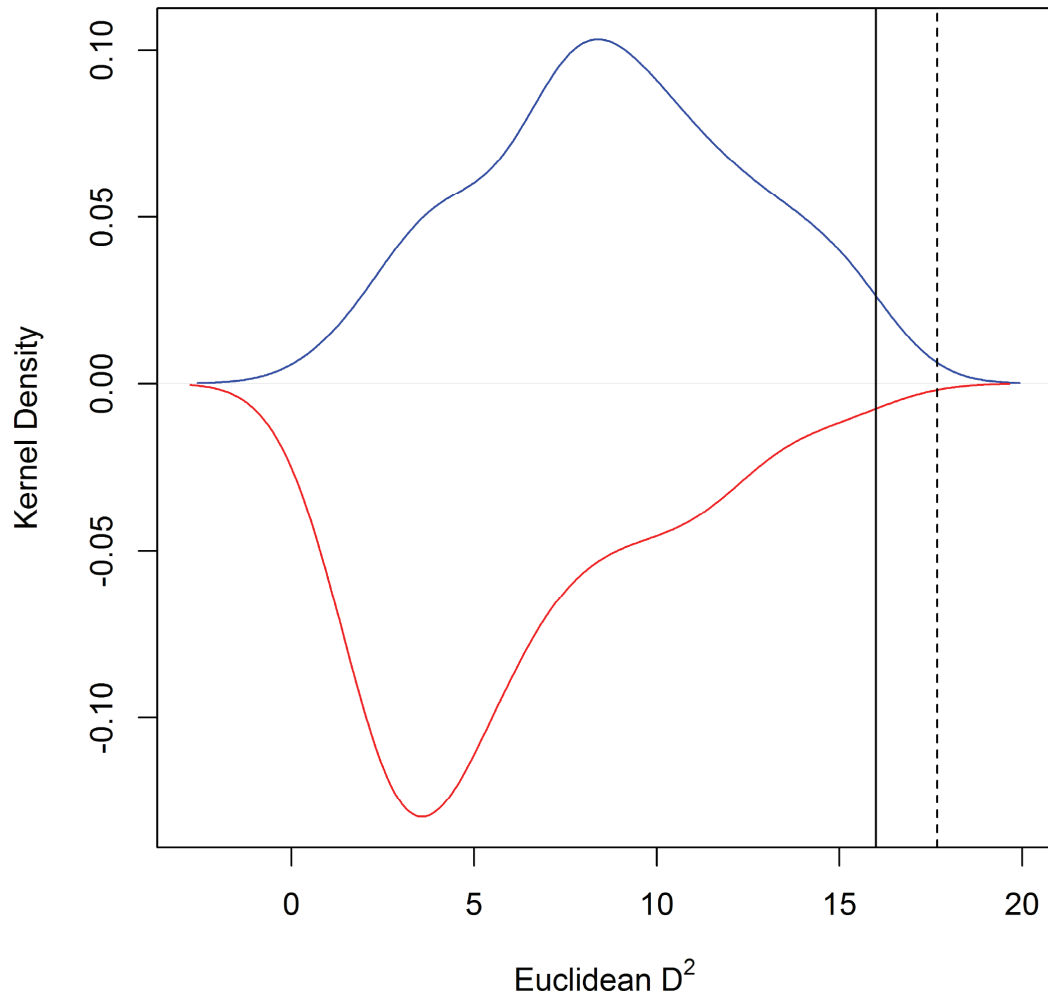


Figure S2. Kernel density mirror plot of Euclidean D^2 analyses of six principal components calculated from 26 cranio-mandibular morphological variables. Related to Figure 1C. The between-species distribution (blue line) was calculated as the distances of all *P. abelii* samples to the *P. pygmaeus* centroid plus all of the *P. pygmaeus* samples to the *P. abelii* centroid, whereas the within-species distribution (red line) was calculated as the distances of all *P. abelii* samples to the *P. abelii* centroid plus all of the *P. pygmaeus* samples to the *P. pygmaeus* centroid. The dotted line represents the distance of the *P.APANULIENSIS* sample to the *P. abelii* centroid (exact permutation test; within-species distribution: p-value<0.001; between-species: p-value<0.001), whereas solid line represents the distance of the *P.APANULIENSIS* samples to the *P. pygmaeus* centroid (within-species: p-value<0.001; between-species: p-value<0.001).

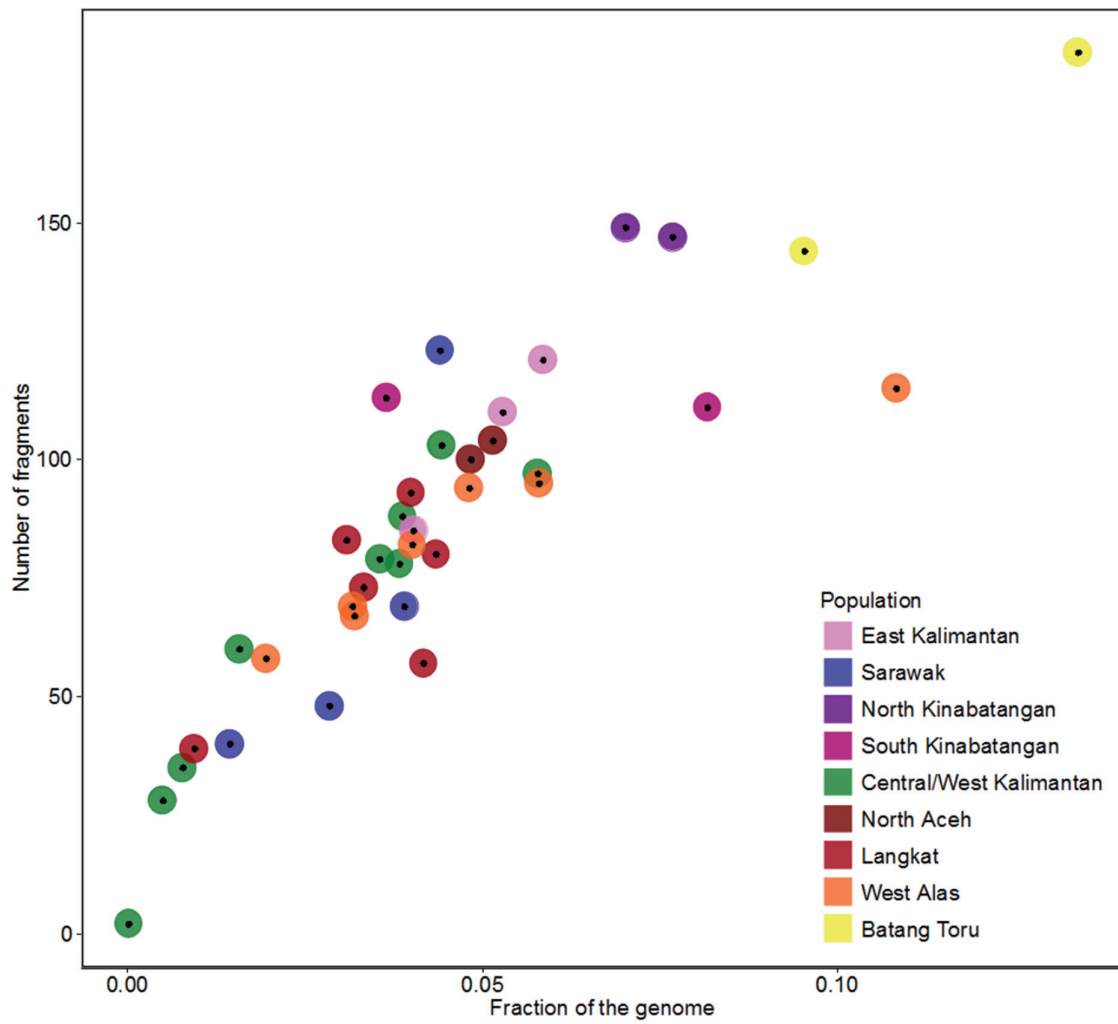


Figure S3. Signatures of recent inbreeding in different orangutan populations. Related to Figure 3C. Number of genomic fragments that are autozygous (y-axis) plotted against the total fraction of the genome covered by such fragments (x-axis). Each dot represents an individual, with sample origins represented by colors corresponding to those in Figure 2A.

Table S1. Summary statistics for the cranio-mandibular variables utilized in this study [mm]. Related to Figure 1B.

Species	PI	PN	NI	PO	LNS	BDS	MBA	MBP	BB	BOE	IB	OB	OH
<i>P. abelii</i>													
Mean	232.05	101.46	136.87	68.53	58.14	138.34	71.39	52.31	103.50	114.44	11.82	36.69	42.54
SD	13.27	10.18	7.10	3.12	4.63	6.79	4.08	4.26	4.32	8.45	1.45	1.74	5.09
Minimum	215.44	85.03	127.93	64.02	49.81	126.59	64.15	45.59	97.00	102.78	8.74	34.65	31.64
1st Quartile	222.07	95.77	130.99	66.31	55.71	134.18	68.89	50.13	101.00	106.96	11.25	35.16	41.50
Median	232.76	102.64	137.88	68.78	58.82	140.50	71.81	52.98	104.25	114.24	12.40	36.68	44.12
3rd Quartile	236.63	107.05	140.41	70.68	60.26	143.51	74.61	54.65	106.63	122.51	12.56	37.74	45.88
Maximum	256.78	116.77	149.32	72.38	64.68	145.59	76.21	58.97	109.00	124.82	13.29	39.48	46.89
n	8	8	8	8	8	8	8	8	8	8	8	8	8
<i>P. pygmaeus</i>													
Mean	234.36	104.80	138.45	64.29	57.58	144.23	71.25	53.64	110.66	115.05	12.23	35.91	41.43
SD	12.10	7.70	8.28	4.68	6.40	8.34	5.12	5.26	6.79	7.41	1.74	2.25	2.85
Minimum	211.58	88.18	120.58	55.50	47.55	128.18	55.69	39.28	98.50	98.01	8.99	29.67	35.29
1st Quartile	227.90	101.97	131.02	60.43	52.07	137.94	68.99	51.27	105.50	111.73	11.22	34.87	39.70
Median	237.86	106.15	138.84	65.09	59.53	146.10	72.25	54.77	111.00	116.28	11.91	35.49	41.66
3rd Quartile	243.66	109.74	146.17	66.83	61.77	148.50	74.43	56.91	114.50	120.32	13.22	36.88	43.32
Maximum	252.40	117.01	150.11	76.10	71.04	158.05	79.20	61.61	125.00	127.82	16.10	40.62	46.03
n	25	25	25	25	21	23	25	25	25	25	25	25	25
<i>P. tapanuliensis</i>													
	224.72	90.80	139.54	69.85	70.52	136.52	65.00	59.94	101.50	120.00	12.42	33.80	33.38
n	1	1	1	1	1	1	1	1	1	1	1	1	1
<i>Permutation tests</i>													
vs. <i>P. abelii</i>	NS	NS	NS	NS	<0.001	NS	NS	NS	NS	NS	NS	NS	NS
vs. <i>P. pygmaeus</i>	NS	NS	NS	NS	0.048	NS	NS	NS	NS	NS	NS	NS	<0.001

PI = Prosthion-Inion Length, PN = Prosthion-Nasion Length, NI = Nasion-Inion Length, PO = Postorbital Breadth, LNS = Nuchal Surface Length, BDS = Nuchal Surface Breadth, MBA = Anterior Muzzle Breadth, MBP = Posterior Muzzle Breadth, BB = Braincase Breadth, BOE = Biorbital Breadth, IB = Interorbital Breadth, OB = Orbital Breadth, OH = Orbital Height.

Table S1 (continued). Summary statistics for the cranio-mandibular variables utilized in this study [mm]. Related to Figure 1B.

Species	DF	PB	PL	ProB	BZAE	ZAT	BT	TMB	TML	PPB	APB	LTTA	TMJA
<i>P. abelii</i>													
Mean	18.19	72.28	91.93	173.43	162.13	9.94	114.95	27.70	34.31	51.65	50.77	33.50	29.83
SD	3.12	3.77	8.70	10.85	6.58	1.50	4.02	2.22	2.75	1.29	3.30	1.19	1.49
Minimum	14.25	66.72	72.80	160.34	152.12	7.57	107.96	23.34	30.00	50.29	47.54	31.71	27.83
1st Quartile	15.41	70.10	91.05	164.82	158.83	9.09	113.07	26.66	32.50	50.63	48.91	32.83	28.43
Median	18.64	71.66	93.33	171.24	161.33	10.23	115.55	27.92	34.83	51.30	50.02	33.29	30.16
3rd Quartile	19.82	74.76	95.51	183.06	164.86	10.90	117.87	29.47	36.34	52.33	51.25	34.10	30.66
Maximum	22.63	78.24	101.29	188.43	174.36	12.06	119.69	30.00	37.77	54.08	58.17	35.24	31.90
n	8	8	8	8	8	8	8	8	8	8	8	8	8
<i>P. pygmaeus</i>													
Mean	14.30	73.59	91.82	171.85	166.19	8.61	119.93	25.67	31.38	49.45	50.08	33.90	31.27
SD	2.75	3.31	6.35	10.88	9.03	1.84	6.09	2.25	3.10	4.50	3.90	2.15	2.65
Minimum	8.39	66.33	80.07	148.42	146.44	3.89	109.33	21.32	25.38	43.87	43.01	28.40	24.68
1st Quartile	12.32	71.57	86.72	163.90	160.72	7.85	115.85	24.19	28.97	47.00	46.57	32.98	30.02
Median	14.75	74.33	92.48	174.43	168.50	8.62	118.83	25.72	31.35	48.20	51.18	34.10	31.78
3rd Quartile	15.70	75.62	96.37	179.31	174.05	9.72	123.81	27.29	33.60	49.90	52.43	35.35	33.12
Maximum	20.58	80.32	103.79	189.95	179.64	12.20	135.28	30.62	38.27	62.39	57.86	37.55	35.40
n	25	25	25	23	25	25	25	22	22	25	25	25.00	25.00
<i>P. tapanuliensis</i>													
	6.04	73.37	82.40	164.30	160.46	10.38	109.48	23.17	29.20	33.78	33.71	23.93	22.46
n	1	1	1	1	1	1	1	1	1	1	1	1	1
Permutation tests													
vs. <i>P. abelii</i>	<0.001	NS	NS	NS	NS	NS	NS	NS	<0.001	<0.001	<0.001	<0.001	<0.001
vs. <i>P. pygmaeus</i>	<0.001	NS	NS	NS	NS	NS	NS	NS	NS	<0.001	<0.001	<0.001	<0.001

DF = Face Depth, PB = Palate Breadth, PL = Palate Length, ProB = Prosthion-Basion Length, BZAE = Bizygomatic Arch Breadth, ZAT = Zygomatic Arch Thickness, BT = Bitympanic Breadth, TMB = Foramen Magnum Breadth, TML = Foramen Magnum Length, PPB = Posterior Pterygoid Breadth, APB = Anterior Pterygoid Breadth, LTTA = Tympanic tube length, TMJA = Temporomandibular joint length.

Table S1 (continued). Summary statistics for the cranio-mandibular variables utilized in this study [mm]. Related to Figure 1B.

Species	BS	PM1M3A	MIB	MM1EB	RA	S	ITT	BiB	HLM	RWA	JIW	JM1EB	JPM1M3A
<i>P. abelii</i>													
Mean	80.41	55.19	39.85	72.01	109.36	61.79	42.11	131.63	159.06	60.55	30.06	59.81	65.99
SD	8.55	2.79	5.22	3.75	5.82	5.53	5.39	3.78	8.92	1.98	1.01	5.03	4.29
Minimum	70.92	50.96	30.12	66.65	102.21	53.71	36.14	127.16	146.64	58.26	28.49	49.03	60.80
1st Quartile	73.86	53.25	38.24	70.05	106.96	58.24	37.31	128.15	152.72	58.96	29.39	58.34	61.95
Median	78.89	55.66	41.21	71.39	107.57	60.92	41.83	131.22	156.98	60.19	30.07	61.08	68.43
3rd Quartile	86.83	57.13	43.22	73.67	110.56	65.51	45.66	134.72	167.33	61.86	30.87	62.35	68.86
Maximum	91.59	58.95	44.68	78.27	121.61	70.02	50.04	137.39	170.34	63.56	31.47	65.45	71.05
n	8	8	7	8	8	8	8	8	8	8	8	8	7
<i>P. pygmaeus</i>													
Mean	82.03	55.33	41.81	73.27	111.44	66.93	42.35	134.04	159.77	65.64	31.96	61.53	69.43
SD	7.32	3.16	2.54	3.45	5.99	5.00	3.25	11.28	12.60	5.22	2.94	3.71	3.26
Minimum	65.34	46.61	34.92	65.73	98.10	60.51	35.98	113.43	116.28	56.28	23.42	52.53	64.58
1st Quartile	78.24	53.99	40.65	71.19	109.16	63.35	40.56	126.02	155.88	63.06	30.60	60.31	67.35
Median	84.85	55.68	41.99	73.35	110.52	65.19	41.90	135.38	161.39	65.25	32.57	61.99	68.72
3rd Quartile	88.18	57.79	43.77	75.48	113.57	70.80	43.68	142.78	166.15	67.96	33.41	64.02	71.65
Maximum	90.93	60.46	45.03	80.25	124.63	79.08	49.32	154.99	180.02	78.90	37.85	66.82	75.85
n	23	24	25	25	20	21	21	21	21	21	21	21	21
<i>P. tapanuliensis</i>													
	77.83	55.27	28.31	62.66	113.61	49.29	31.80	119.98	150.58	55.94	24.44	55.32	70.00
n	1	1	1	1	1	1	1	1	1	1	1	1	1
Permutation tests													
vs. <i>P. abelii</i>	NS	NS	<0.001	<0.001	NS	<0.001	<0.001	<0.001	NS	<0.001	<0.001	NS	NS
vs. <i>P. pygmaeus</i>	NS	NS	<0.001	<0.001	NS	<0.001	<0.001	NS	NS	NS	0.048	NS	NS

BS = Basion-Staphylion Length, PM1M3A = Maxillary Length of PM1-M3, MIB = Maxillary Incisor Complex Breadth, MM1EB = External Breadth of the Maxilla at M1, RA = Ramus Height, S = Symphysis Length, ITT = Inferior transverse torus, BiB = Bicondylar Breadth, HLM = Horizontal Length, RWA = Ramus Width, JIW = Mandibular Incisor Complex Breadth, JM1EB = External Breadth of the Mandible at M1, JPM1M3A = Mandibular Length of PM1-M3.

Table S2. Summary statistics for the dental variables utilized in this study [mm]. Related to Figure 1B.

Species	UCB	LCB	LM1L	LM1B	LM1A
<i>P. abelii</i>					
Mean	17.90	15.96	13.12	10.81	141.86
SD	1.77	0.96	0.57	0.60	10.23
Minimum	15.67	14.34	12.48	10.08	128.51
1st Quartile	16.76	15.61	12.66	10.36	133.35
Median	17.37	16.05	13.00	10.87	145.43
3rd Quartile	19.38	16.29	13.56	11.18	148.32
Maximum	20.54	17.55	13.89	11.68	155.74
n	8	8	7	7	7
<i>P. pygmaeus</i>					
Mean	18.08	17.03	13.46	11.22	151.04
SD	1.57	1.61	0.78	0.70	13.58
Minimum	14.82	14.46	11.38	10.11	126.17
1st Quartile	17.37	15.59	13.17	10.57	140.12
Median	17.85	17.20	13.50	11.31	147.79
3rd Quartile	19.27	18.27	13.83	11.74	162.36
Maximum	20.86	19.60	15.01	12.45	171.56
n	19	19	20	20	20
<i>P. p. palaeosumatrensis</i>					
Mean	20.94	17.28	14.99	13.05	195.71
SD	1.91	1.47	0.53	0.58	14.09
Minimum	18.30	15.30	14.00	12.10	175.45
1st Quartile	19.10	16.05	14.60	12.70	183.80
Median	21.20	17.00	14.90	13.00	193.50
3rd Quartile	22.00	18.15	15.40	13.48	205.74
Maximum	24.60	20.50	16.20	14.50	234.90
n	21	39	90	90	90
<i>P. tapanuliensis</i>					
	21.50	19.44	13.65	11.37	155.20
n	1	1	1	1	1
<i>Permutation tests</i>					
vs. <i>P. abelii</i>	<0.001	<0.001	NS	NS	NS
vs. <i>P. pygmaeus</i>	NS	NS	NS	NS	NS
vs. <i>P. p. palaeosumatrensis</i>	NS	NS	<0.001	<0.001	<0.001

UCB = Upper canine breadth, LCB = Lower canine breadth, LM1L = Lower M1 length, LM1B = Lower M1 breadth, LM1A = Lower M1 area.

Table S3. Summary statistics for the 15 long call variables utilized in this study. Related to STAR Methods.

Species	No. of pulses	Call Dur [s]	Sound Dur [s]	Interval Dur [s]	Max Freq R [Hz]
<i>P. abelii</i>					
Mean	40.74	72.70	0.61	1.09	558.83
SD	9.63	24.17	0.08	0.19	121.73
Minimum	26.50	46.22	0.47	0.76	369.76
1st Quartile	32.94	50.42	0.57	0.98	468.26
Median	38.75	65.20	0.61	1.12	557.78
3rd Quartile	47.67	96.25	0.67	1.22	642.11
Maximum	56.50	113.60	0.74	1.46	746.86
n	14	14	14	14	14
<i>P. pygmaeus</i>					
Mean	25.41	53.59	0.69	1.37	706.99
SD	7.72	13.73	0.18	0.34	184.11
Minimum	10.00	28.76	0.43	0.80	257.25
1st Quartile	21.00	45.79	0.57	1.06	621.98
Median	25.00	51.80	0.66	1.39	689.88
3rd Quartile	29.00	60.68	0.79	1.63	836.52
Maximum	45.00	89.36	1.28	1.97	998.74
n	29	29	29	29	27
<i>P. tapanuliensis</i>					
Mean	57.11	112.06	0.66	1.06	830.64
SD	5.97	0.39	0.04	0.06	42.15
Minimum	52.89	111.78	0.63	1.02	800.84
1st Quartile	55.00	111.92	0.64	1.04	815.74
Median	57.11	112.06	0.66	1.06	830.64
3rd Quartile	59.22	112.19	0.67	1.08	845.55
Maximum	61.33	112.33	0.68	1.10	860.45
n	2	2	2	2	2
<i>Permutation tests</i>					
vs. <i>P. abelii</i>	NS	NS	NS	NS	<0.001
vs. <i>P. pygmaeus</i>	<0.001	NS	NS	NS	NS

No. of pulses = Number of pulses, Call Dur = Duration of call, Sound Dur = Duration of sound, Interval Dur = Duration of interval, Max Freq R = Maximum frequency of roar (R) pulse type.

Table S3 (continued). Summary statistics for the 15 long call variables utilized in this study. Related to STAR Methods.

Species	Min Freq R [Hz]	Peak Freq R [Hz]	Shape R [Hz/s]	Freq Max [Hz]	Freq Min [Hz]
<i>P. abelii</i>					
Mean	141.77	310.61	709.07	824.29	64.04
SD	39.16	60.44	155.29	193.91	30.40
Minimum	88.90	186.82	450.06	460.38	17.64
1st Quartile	103.36	279.97	572.28	732.78	49.39
Median	148.70	294.25	739.86	837.01	61.87
3rd Quartile	173.99	362.27	833.23	948.13	75.76
Maximum	200.53	400.52	934.08	1111.25	145.50
n	14	14	14	14	14
<i>P. pygmaeus</i>					
Mean	177.36	403.82	749.46	984.66	62.13
SD	61.70	111.90	247.91	291.69	29.46
Minimum	74.08	202.17	230.78	354.29	10.58
1st Quartile	135.31	336.22	642.39	896.06	45.86
Median	173.87	387.60	730.15	977.19	57.00
3rd Quartile	215.93	436.23	870.72	1167.10	77.16
Maximum	361.07	732.13	1372.05	1498.60	144.44
n	27	27	27	29	29
<i>P. tapanuliensis</i>					
Mean	199.17	399.56	1036.53	1136.15	87.69
SD	7.57	19.16	118.19	128.95	10.08
Minimum	193.82	386.02	952.96	1044.97	80.57
1st Quartile	196.50	392.79	994.74	1090.56	84.13
Median	199.17	399.56	1036.53	1136.15	87.69
3rd Quartile	201.85	406.33	1078.31	1181.74	91.26
Maximum	204.53	413.11	1120.10	1227.33	94.82
n	2	2	2	2	2
Permutation tests					
vs. <i>P. abelii</i>	NS	NS	<0.001	NS	NS
vs. <i>P. pygmaeus</i>	NS	NS	NS	NS	NS

Min Freq R = Minimum frequency of roar (R) pulse type, Peak Freq R = Peak frequency of roar pulse type, Shape R = Average shape of roar pulse type, Freq Max = Maximum frequency of call, Freq Min = Minimum frequency of call.

Table S3 (continued). Summary statistics for the 15 long call variables utilized in this study. Related to STAR Methods.

Species	Rate [pulses/20s]	Huitus [%]	Roar [%]	Sigh [%]	Intermediary [%]
<i>P. abelii</i>					
Mean	0.81	10.26	54.57	6.54	5.31
SD	0.11	13.68	15.66	4.29	5.41
Minimum	0.62	0.00	19.35	0.00	0.00
1st Quartile	0.72	3.15	48.03	5.44	1.10
Median	0.81	5.61	53.85	6.84	4.83
3rd Quartile	0.89	8.68	66.53	8.23	6.96
Maximum	0.97	48.39	75.76	13.51	16.67
n	14	14	14	14	14
<i>P. pygmaeus</i>					
Mean	0.52	16.26	28.36	15.51	11.02
SD	0.13	15.58	17.23	18.17	9.26
Minimum	0.30	0.00	0.00	0.00	0.00
1st Quartile	0.45	0.00	20.29	4.35	4.35
Median	0.48	16.54	26.92	8.00	8.21
3rd Quartile	0.64	23.11	35.55	20.30	15.38
Maximum	0.79	64.00	80.95	80.00	41.67
n	29	29	29	29	29
<i>P. tapanuliensis</i>					
Mean	0.88	7.80	39.58	20.47	1.98
SD	0.08	11.03	0.81	10.24	2.80
Minimum	0.82	0.00	39.01	13.23	0.00
1st Quartile	0.85	3.90	39.29	16.85	0.99
Median	0.88	7.80	39.58	20.47	1.98
3rd Quartile	0.91	11.69	39.87	24.09	2.97
Maximum	0.93	15.59	40.15	27.71	3.96
n	2	2	2	2	2
<i>Permutation tests</i>					
vs. <i>P. abelii</i>	NS	NS	NS	<0.001	NS
vs. <i>P. pygmaeus</i>	<0.001	NS	NS	NS	NS

Rate = Number of pulses per 20 s, Huitus = Percent number of huitus (H) pulse type, Roar = Percent number of roar (R) pulse type, Sigh = Percent number of sigh (S) pulse type, Intermediary = Percent number of intermediary (I) pulse type.

Table S4. Details of study individuals. Related to Figure 2A.

Species	Sampling area	Individual ID	Name	Sex	Depth ^a	Source	Comments and origin details, if available
<i>P. abelii</i>	Langkat	PA_KB4661	Bubbles	M	4.76	[S1]	Wild-born
<i>P. abelii</i>	Langkat	PA_KB5883	Sibu	M	4.99	[S1]	Wild-born
<i>P. abelii</i>	Langkat	PA_A947	Elsi	F	27.39	[S2]	Wild-born
<i>P. abelii</i>	Langkat	PA_A948	Kiki	F	23.71	[S2]	Wild-born
<i>P. abelii</i>	Langkat	PA_A950	Babu	F	26.28	[S2]	Wild-born
<i>P. abelii</i>	Langkat	PA_A952	Buschi	M	21.03	[S2]	Wild-born
<i>P. abelii</i>	North Aceh	PA_A949	Dunja	F	27.39	[S2]	1 st Generation by 456 and 457 both wild-born Sumatra
<i>P. abelii</i>	North Aceh	PA_B018	Jeff	M	16.31	This study	Wild-born; Desa Seuneubok Bayu, Indra Makmu district
<i>P. abelii</i>	West Alas	PA_KB4361	Likoe	F	5.66	[S1]	Wild-born
<i>P. abelii</i>	West Alas	PA_SB550	Doris	F	4.86	[S1]	Wild-born
<i>P. abelii</i>	West Alas	PA_B017	Miky	F	13.74	This study	Wild-born; Aluebillie, Aceh Nagan Raya, Aceh province
<i>P. abelii</i>	West Alas	PA_A953	Vicky	F	17.78	This study	Wild-born
<i>P. abelii</i>	West Alas	PA_A955	Suma	F	25.27	This study	Wild-born
<i>P. abelii</i>	West Alas	PA_A964	Rochelle	F	11.06	This study	Wild-born
<i>P. abelii</i>	West Alas	PA_B020	Maini	F	16.3	This study	Wild-born; Aceh Sealatan near Suaq Balimbing
<i>P. tapanuliensis</i>	Batang Toru	PA_KB9258	Baldy	F	5.79	[S1]	Wild-born
<i>P. tapanuliensis</i>	Batang Toru	PA_KB9258	Baldy	F	5.79	[S1]	Wild-born

^amean effective whole-genome sequencing coverage (estimated from the quality filtered BAM files).

Table S4 (continued). Details of study individuals. Related to Figure 2A.

Species	Sampling area	Individual ID	Name	Sex	Depth ^a	Source	Comments and origin details, if available
<i>P. pygmaeus</i>	Central Kalimantan	PP_KB4204	Dolly	M	5.61	[S1]	Wild-born
<i>P. pygmaeus</i>	Central Kalimantan	PP_KB5404	Billy	F	12.24	[S1]	Wild-born
<i>P. pygmaeus</i>	Central Kalimantan	PP_KB5405	Dennis	M	5.61	[S1]	Wild-born
<i>P. pygmaeus</i>	Central Kalimantan	PP_A940	Temmy	F	21.8	[S2]	1 st Generation by 793 and 794 both wild-born Borneo
<i>P. pygmaeus</i>	Central Kalimantan	PP_A941	Sari	F	23.17	[S2]	1. Gen. by 202 and 322 both wild-born Borneo
<i>P. pygmaeus</i>	Central Kalimantan	PP_A943	Tilda	F	24.17	[S2]	Wild-born
<i>P. pygmaeus</i>	Central Kalimantan	PP_A944	Napoleon	M	23.32	[S2]	Wild-born
<i>P. pygmaeus</i>	Central Kalimantan	PP_A938	Lotti	F	18.62	This study	1 st Generation by 358 and 422 both wild-born Borneo
<i>P. pygmaeus</i>	West Kalimantan	PP_A983	Claus	M	29.71	This study	Wild-born; Pontianak
<i>P. pygmaeus</i>	East Kalimantan	PP_KB5543	Louis	M	6.03	[S1]	Wild-born
<i>P. pygmaeus</i>	East Kalimantan	PP_A984	Barong	F	29.89	This study	Wild-born; Taman Nasional Kutai
<i>P. pygmaeus</i>	East Kalimantan	PP_A985	Panjul	M	30.13	This study	Wild-born; Taman Nasional Kutai
<i>P. pygmaeus</i>	North Kinabatangan	PP_A987	Tara	F	30.65	This study	Wild-born; Bukit Garam, Kinabatangan area
<i>P. pygmaeus</i>	North Kinabatangan	PP_A988	Kala	M	31.06	This study	Wild-born; Kg. Tikolod, Tambunan
<i>P. pygmaeus</i>	South Kinabatangan	PP_5062	Ampal	M	13.81	This study	Wild-born; Lahad Datu, Kinabatangan area
<i>P. pygmaeus</i>	South Kinabatangan	PP_A989	Micelle	F	27.3	This study	Wild-born; Lahad Datu, Kinabatangan area
<i>P. pygmaeus</i>	Sarawak	PP_KB5406	Dinah	F	4.9	[S1]	Wild-born
<i>P. pygmaeus</i>	Sarawak	PP_A939	Nonja	F	20.48	[S2]	1 st Generation by 1052 and 1012 both from Sarawak
<i>P. pygmaeus</i>	Sarawak	PP_A942	Gusti	F	23.12	This study	1 st Generation by 1435 and 1392 both wild-born Borneo
<i>P. pygmaeus</i>	Sarawak	PP_A946	Kajan	M	22.39	This study	Wild-born

^amean effective whole-genome sequencing coverage (estimated from the quality filtered BAM files).

Table S5. Parameter estimation of the best supported models in the ABC and G-PhoCS analyses. Related to Figure 3B.

<i>ABC</i>				
Parameter ^a	Prior distribution	Mode	Mean	95%-HPD ^b
N _{NOW} BO (4)	loguniform (300–32,000)	1,487	1,759	407–8,002
N _{NOW} NT (2)	loguniform (300–32,000)	2,854	3,212	517–21,691
N _{STRUC} NT (2)	loguniform (3,000–320,000)	19,925	26,795	3,736–197,419
N _{NOW} ST	loguniform (300–32,000)	2,520	2,429	524–10756
N _{ANC} ST	loguniform (1,000–100,000)	35,874	28,907	7,522–99,885
N _{BN} BO	loguniform (300–32,000)	4,473	3,719	523–27,948
N _{ANC} BO	loguniform (3,000–320,000)	30,655	36,257	5,924–266,244
N _{ANC} NT	loguniform (1,000–100,000)	53,811	29,654	5,115–99,885
T _{BNEND} BO	uniform (8,750–400,000)	71,969	125,689	8,848–272,775
T _{BNDUR} BO	uniform (250–100,000)	33,583	46,508	924–92,087
T _{SPLIT} BO	uniform (400,000–1,500,000)	674,055	681,760	427,878–921,400
T _{SPLIT} NT	uniform (1,500,000–4,000,000)	3,382,200	2,827,150	1,712,005–3,977,250
T _{DEC} NT	uniform (250–100,000)	82,635	54,372	10,126–99,975
T _{STRUC} NT	uniform (100,000–1,500,000)	1,057,388	873,195	241,301–1,499,650
T _{MIGSTOP}	uniform (8,750–400,000)	303,118	253,968	82,680–399,903
NmWBO	loguniform (0.030–32.000)	6.818	1.272	0.060–31.568
NmWNT	loguniform (0.030–32.000)	0.128	0.594	0.032–14.973
NmBOST	loguniform (0.003–3.200)	0.016	0.021	0.003–0.127
NmSTBO	loguniform (0.003–3.200)	0.003	0.007	0.003–0.021
NmNTST	loguniform (0.010–10.000)	0.294	0.228	0.019–2.116
NmSTNT	loguniform (0.010–10.000)	0.86	0.687	0.058–9.166

^a, BO = Borneo, NT = Sumatra north of Lake Toba, ST = Sumatra south of Lake Toba, N_{NOW} = current effective population size (N_e), N_{BN} = N_e during population bottleneck, N_{ANC} = ancestral N_e, N_{STRUC} = N_e before recent decline (number of populations of this size), T_{BNEND} = time since population bottleneck ended, T_{BNDUR} = duration of bottleneck, T_{SPLIT} = population split time, T_{DEC} = time since population decline, T_{STRUC} = time since establishment of population structure, T_{MIGSTOP} = time since migration between BO and ST stopped (all times were converted to years assuming a generation time of 25 years), NmWBO = number of migrants per generation among populations on Borneo, NmWNT = number of migrants among populations north of Lake Toba, NmXY = number of migrants in X from Y; ^b, 95%-highest posterior density interval.

Table S5 (continued). Parameter estimation of the best supported models in the ABC and G-PhoCS analyses. Related to Figure 3B.

<i>G-PhoCS</i>				
Parameter ^a	Prior distribution ^b	Mode	Mean	95%-HPD ^c
N _{NOW} BO	Gamma ($\alpha=1$; $\beta=500$)	17,939	17,992	17,655–18,338
N _{NOW} NT	Gamma ($\alpha=1$; $\beta=500$)	16,123	16,114	15,588–16,655
N _{NOW} ST	Gamma ($\alpha=1$; $\beta=500$)	26,787	26,791	26,113–27,477
N _{ANC} BOST	Gamma ($\alpha=1$; $\beta=500$)	114,303	114,451	110,626–118,704
N _{ANC} PONGO	Gamma ($\alpha=1$; $\beta=500$)	33,162	33,223	32,316–34,119
T _{SPLIT} BOST	Gamma ($\alpha=1$; $\beta=2000$)	575,551	578,150	563,217–593,200
T _{SPLIT} PONGO	Gamma ($\alpha=1$; $\beta=500$)	2,273,045	2,278,133	2,208,383–2,351,917
m_BO->ST	Gamma ($\alpha=0.002$; $\beta=0.00001$)	4.45×10^{-6}	4.45×10^{-6}	$4.08\text{--}4.80 \times 10^{-6}$
m_ST->BO	Gamma ($\alpha=0.002$; $\beta=0.00001$)	1.17×10^{-6}	1.20×10^{-6}	$0.95\text{--}1.46 \times 10^{-6}$
m_NT->ST	Gamma ($\alpha=0.002$; $\beta=0.00001$)	3.19×10^{-6}	3.27×10^{-6}	$2.55\text{--}3.94 \times 10^{-6}$
m_ST->NT	Gamma ($\alpha=0.002$; $\beta=0.00001$)	8.28×10^{-5}	8.29×10^{-5}	$7.98\text{--}8.60 \times 10^{-5}$
m_BOST->NT	Gamma ($\alpha=0.002$; $\beta=0.00001$)	8.39×10^{-5}	8.53×10^{-5}	$5.47\text{--}11.44 \times 10^{-5}$
m_NT->BOST	Gamma ($\alpha=0.002$; $\beta=0.00001$)	6.87×10^{-12}	2.18×10^{-10}	$0.0015\text{--}11.73 \times 10^{-10}$

^a, BO = Borneo, NT = Sumatra north of Lake Toba, ST = Sumatra south of Lake Toba, BOST = ancestral population of BO and ST, PONGO = ancestral population of all orangutans, N_{NOW} = current effective population size, N_{ANC} = ancestral effective population size, T_{DIV} = population split time in years, m_X->Y = migration rate per generation from X to Y forward in time; ^b, prior distribution of mutation-scaled parameters; ^c, 95%-highest posterior density interval. All scaled estimates from G-PhoCS were converted to absolute values assuming a mutation rate of 1.5×10^{-8} mutations per base pair per generation and a generation time of 25 years.

Table S6. PCR primers for Sanger sequencing of mitogenomes. Related to STAR Methods.

Primer name	Primer sequence (3'-5')	Primer position^a
F1	GYTTGGTCCTRGCTTTC	77
R1	AGTACRCTTACCATGTTAC	1004
F2	ACACACCGCCCGTCAC	902
R2	CAGGTCAATTTCACTGGT	2109
F3	CATCACCTCTAGCATTAC	1931
R3	ATTAGGGCGTAGTTWGAG	3120
F4	AAGATGGCAGAGCCCG	2658
R4	CAACATTTTCGGGGTATG	3874
F5	CTGACRAAAGAGTTACTTTG	3698
R5	GGGCTTAGCTTAATTAAG	5076
F6	CCAAGAGCCTTCAAAGC	4958
R6	CYGTRAATATRTGGTGGGC	6224
F7	TWCTCYCACCCAGGAGC	5732
R7	GGGGYTGGCTTGAAACC	6917
F8	AAAGGAAGGAATCGAACC	6873
R8	GTCTTTAACTTAAAAGGTAA	7776
F9	GAGGCCCAYTGCAAAGC	7729
R9	TGGTGGCCTTGGTATGT	8858
F10	CYACCCARCTWTCCATAAA	8250
R10	CCTCATCAGTAGATGGAG	9425
F11	TTCCACGGCCTCCACG	9253
R11	GATAAGGGGTCGGAGG	10384
F12	AAAYAAATGATTTGCACTCAT	9863
R12	AAGCTTCAGGGGGTTTG	11125
F13	CGACAAACAGAYCTAAAATC	11047
R13	GTTGATRTTTGGGTCTGAG	12135
F14	GTGCAACTCCAAATAAAAG	11770
R14	AGGGCTCAGGCGTTGG	13016
F15	TCTGCACCCAYGCCTTC	12776
R15	GTATGATGGTTGTTTTTGG	13943
F16	GCACCCGCACCAATAG	13687
R16	GGCCTCAYGGGAGGAC	14609
F17	CGAGAYGTAAACTACGGC	14411
R17	AGTTAAGTRCTTTTTCTCTG	15435
F18	CAAGCAACAGAGCATAAC	15130
R18	TGTCTTATTTAAGGGGAAC	16017
F19	CTGTATCCGGCATCTGG	15943
R19	CGCGGTGGCTGGCAC	324

^a, Sequence positions (3'-end) on the *Pongo abelii* reference mitochondrial genome NC_002083.

Supplemental References

- S1. Locke, D.P., Hillier, L.W., Warren, W.C., Worley, K.C., Nazareth, L.V., Muzny, D.M., Yang, S.-P., Wang, Z., Chinwalla, A.T., Minx, P., et al. (2011). Comparative and demographic analysis of orang-utan genomes. *Nature* 469, 529-533.
- S2. Prado-Martinez, J., Sudmant, P.H., Kidd, J.M., Li, H., Kelley, J.L., Lorente-Galdos, B., Veeramah, K.R., Woerner, A.E., O'Connor, T.D., Santpere, G., et al. (2013). Great ape genetic diversity and population history. *Nature* 499, 471-475.